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ABSTRACT

Early life exposure to local financial institutions increases household financial inclusion and leads to long-term improvements in consumer credit outcomes. We identify the effect of local financial markets using Congressional legislation that led to unintended differences in financial market development across Native American reservations. Individuals from financially underdeveloped reservations enter consumer credit markets later, and upon reaching adulthood, have ten point lower credit scores and four percentage point more delinquent accounts. These effects are long-lived and depreciate slowly after individuals move to more developed areas. Formative exposures to local banking improve consumer credit behavior by increasing financial literacy and financial trust.

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1. Introduction

Consumer credit markets have high stakes and affect most households. Previous research identifies several important determinants of attitude formation and risk-taking in financial markets, and recent research on credit market behavior shows how individual characteristics, such as cognitive biases and financial education (Agarwal et al., 2009; Keys and Wang, 2018; Brown et al., 2016), affect financial health outcomes, including participation rates and debt repayment.¹ Nonetheless, a large component of

¹ Personal attributes that affect financial risk-taking include socioeconomic status (Kuhnen and Miu, 2017), trust (Gurun et al., 2017; Gianetti and Wang, 2016), ambiguity aversion (Dimmock et al., 2016), physical stature (Cronqvist et al., 2016; Addoum et al., 2016), and genetics (Barnea et al., 2010; Grinblatt et al., 2012; Cronqvist and Siegel, 2015). Anagol et al. (2015) and Knüpfer et al. (2017) consider the role of previously experienced financial returns.

the cross-sectional variation in household credit behavior remains unexplained.

This paper identifies an important, but unappreciated, determinant of household financial health—formative exposure to local financial institutions. Beyond the key role early life experiences play in the formation of financial attitudes and beliefs (e.g., Malmendier and Nagel, 2011), our work shows that the financial markets individuals encounter at a young age have a large, persistent impact on how they build and manage credit over their lifetime. We attribute these long-run effects to greater bank-related financial literacy and more trust in financial institutions that come from having formative exposure to banking.

There are several challenges in identifying how local financial markets affect household financial behavior and credibly evaluating the mechanisms behind these effects. One key challenge is distinguishing the effects of local financial development from differences in economic opportunity, particularly given the lack of exogenous variation in financial development (e.g., Levine, 2005). Even regulatory changes that affect lending activity are not random and often coincide with other economic factors that influence the supply and demand for credit.² Furthermore, it is difficult to construct micro-level measures of consumer credit decisions that are geographically precise, comparable across households, and informative about the factors underlying these financial behaviors. We confront these empirical challenges using panel data on individual-level credit histories from the Federal Reserve Bank of New York Consumer Credit Panel (FRBNY-CCP) — a 5% sample of Equifax consumer credit records — to examine how exposure to financial markets at a young age affects individual engagement with credit markets and long-run financial health. Our tests utilize stark long-run differences in financial development across Native American reservations attributable to US Congressional legislation called Public Law 280 (PL280), passed in 1953. PL280 imposed state court enforcement of debt contracts on a subset of reservations, leaving tribal courts in place for the remaining (non-PL280) reservations (Anderson and Parker, 2008). Although economic and financial conditions were similar on state and tribal court reservations prior to PL280 (Anderson and Parker, 2017), in the decades following the law's passage, banking activity expanded more rapidly on state court reservations. Brown et al. (2017) show that tribal court reservations have approximately 20% fewer bank branches per capita by the 2000s, which they attribute to the greater predictability of debt contract enforcement under state courts.

This particular episode of law-driven financial development is especially attractive for our purposes because PL280 does not directly affect the enforcement of most consumer credit contracts. The law mainly applies to secured lending, and most consumer credit contracts (e.g., credit cards) are unsecured with the lender's contract

indicating the location where any disputes are adjudicated.³ Thus, from the standpoint of consumer credit, the most relevant difference arising from PL280 is that young borrowers on state court reservations had greater exposure to local banking than did their counterparts on tribal court reservations.

We find that individuals who grow up on tribal court reservations are 20 percentage points less likely to have a credit report compared to individuals on state court reservations. And for those who eventually develop a credit history, the rate of entry into formal credit markets is around 10% slower for the individuals growing up in tribal court areas. These differences in formative engagement with consumer credit markets have economically meaningful long-run consequences for household financial health, measured by both credit riskscores and delinquency rates. After controlling for area incomes, young borrowers growing up on tribal court reservations have seven to ten percentage point lower credit scores and two to four percentage point higher delinquency rates. The independent effect of growing up on a tribal court reservation on credit scores is similar to the effect of reducing annual personal incomes by \$6000. These lower credit scores of tribal court residents also significantly impede access to finance, and make financial products more expensive conditional on obtaining access. For example, using our estimates for 30-year-old individuals, six percentage point fewer tribal court residents would qualify for a conventional rate mortgage (have a credit score above 620). On the extensive margin, borrowers from tribal court reservations who do qualify have a 5% higher cost of financing a mortgage than their state court counterparts.

Digging deeper, we present several tests that shed light on the mechanisms through which formative exposure to banking leads to improvements in consumer credit profiles. First, we examine the role of financial literacy, using variation from state-level reforms to financial, economics, and mathematics education that were studied in Brown et al. (2016). We find that mandated financial literacy training substantially mitigates the negative effect that growing up without finance has on consumer engagement with credit markets, whereas analogous general economics and math education mandates do not have the same offsetting effects. Because these education mandates are unrelated to PL280 assignment, this evidence not only suggests that financial education can substitute for early life exposure to finance but also that general differences in education or income do not fully explain the adverse consequences of growing up without finance. Second, we examine whether growing up without finance affects consumers' understanding of credit market information using data from the Consumer Financial Protection Bureau's (CFPB) consumer complaints database. The complaints data show that individuals who grow up without finance are significantly more likely to file credit information complaints that receive no relief. Consistent with the

² For example, the CARD Act of 2009 was enacted precisely because of problems with how individual consumers used credit cards, but nonetheless had sweeping effects on consumer financial health (e.g., Agarwal et al., 2015; Debbaut et al., 2016).

³ In addition, mortgages on reservations have full guarantees provided by the Federal Housing Association because tribal lands — whether on a tribal or state court reservation — are not allowed to be used as collateral.

evidence on financial literacy mandates, these findings suggest individuals who grow up without finance have less understanding of credit market information. Third, we commissioned Qualtrics to sample a broad cross-section of Native Americans on their experiences with financial institutions, attitudes on financial matters, and skills at solving basic consumer financial problems. Controlling for individual-specific levels of income and education, survey respondents who grew up near a bank have better bank-related financial literacy (e.g., understanding compound interest) and more trust in financial institutions. We find no indication that formative exposure to banking operates through other notable mechanisms. Specifically, growing up near a bank is unrelated to general education attainment and literacy on financial topics unrelated to retail banking services (stocks and bonds) and is unrelated to overall generic (nonfinance) levels of trust.

A natural competing explanation is that individuals with more early life exposure to finance also have systematically better economic opportunities. Indeed, [Brown et al. \(2017\)](#) show that the differences in financial development across state and tribal court reservations led to differences in reservation area economic development. Although these differences in economic development undoubtedly affect consumer credit profiles, we present several complementary analyses that imply the effects of growing up without finance extend beyond differences in economic opportunity or credit supply decisions. First, we find similar results for delinquencies as we do for credit scores. Because delinquencies condition on being supplied credit in the first place, our findings suggest that early exposure to finance affects the way individuals manage credit, and not merely their ability to obtain it. Second, our specifications using credit bureau data control for census tract income, while the results using survey data directly control for individual-level personal incomes. We also show robustness to controlling for other observable economic and demographic factors, such as employment rates and education. Third, to the extent that we cannot fully control for all relevant factors, we provide evidence that state-level reforms uncorrelated with either the PL280 treatment or reservation area economic opportunities mitigate the effects of growing up without finance. Specifically, the effects are attenuated in states with mandated financial education training and in states affected by the Interstate Bank Branching Efficiency Act of 1994 (IBBEA), which expanded bank branching into low-income, rural areas without significantly affecting demand conditions ([Celierier and Matray, 2016](#)). Fourth, the survey shows that growing up near banks affects bank-centric financial literacy and trust in financial institutions but is unrelated to either stock/bond literacy or generalized trust, suggesting that growing up near banks does not solely reflect better economic opportunities.

We present several additional tests that indicate the differences in consumer credit profiles are due to early life experiences, rather than stronger current or future economic opportunities. In particular, we find that PL280 does not lead to differences in the financial health of individuals born before the passage of the law, a time when there were no differences in financial conditions across tribal

and state court reservations. Because these individuals and the young borrowers we study in our main tests are exposed to the same current economic conditions, this test suggests that our findings capture the effects of formative experiences on financial health. In addition, we follow the financial health of young borrowers over time when they leave reservations (e.g., [Guiso et al. \(2004b\)](#) use a similar approach to disentangle the effect of formative exposure from current conditions). Our tests compare state and tribal court consumers who both move to the same off-reservation location. Hence, these consumers presumably have similar economic opportunities in their new location. Moving has a relatively stronger effect on financial health for individuals leaving tribal court reservations, but it takes many years of exposure to off-reservation areas for financial health to converge, even for tribal and state court individuals who move to the same area.

These findings provide some of the first causal evidence linking the local provision of finance across institutional environments with consumer financial health. This evidence offers a new perspective on the real consequences of financial development (e.g., [King and Levine, 1993](#); [Levine et al., 2000](#)).⁴ Our work shows that household credit outcomes benefit from financial market development, most notably via higher credit scores and lower delinquency rates. Moreover, our findings suggest these consumer-side benefits are not just the result of better direct access to local bank loans: formative exposure to financial markets improves financial literacy, trust in financial institutions, and has a positive impact on the way consumers build and manage credit, benefits of financial development not emphasized in prior work.

Our study is particularly relevant for the strand of the financial development literature that focuses on the economic effects of stronger local financial markets (e.g., [Javaratne and Strahan \(1996\)](#); [Guiso et al. \(2004a\)](#); [Butler et al. \(2016\)](#)). Notably, several studies show that better access to financial intermediaries improves financial inclusion, typically measured as the formal use of banking services (e.g., [Beck et al. \(2007\)](#); [Demirguc-Kunt and Klapper \(2013\)](#)). Relative to this work, our results show that the effects of strong local financial markets extend well beyond the use of standard banking services, influencing how individuals use and access revolving credit accounts and how their credit profiles evolve over time. In a similar vein, our study adds to previous evidence on the benefits of a more developed local financial market for young and small firms (e.g., [Rice and Strahan \(2010\)](#); [Krishnan et al. \(2015\)](#)) by documenting long-term effects on the financial behaviors of young individuals. Having a more complete accounting of the effects of local financial markets is particularly important given the shift toward consolidation and the nationalization of financial activity in recent years (e.g., [Goetz et al., 2016](#)).⁵

⁴ Much of the research on consumer credit focuses on the pros and cons of access to high interest rate loans, particularly for low-income borrowers (e.g., [Karlan and Zinman, 2010](#); [Melzer, 2011](#)).

⁵ The changing nature of local financial activity has also been widely discussed in the popular press. For example, see Minyoung Park, "America's brick-and-mortar banks are vanishing," Yahoo Finance, June 22, 2016.

Our work is also part of a growing literature that uses credible identification and within-country variation to evaluate the economic effects of institutions, financial markets, and legal rules (e.g., [Butler and Cornaggia \(2011\)](#); [Berkowitz et al. \(2015\)](#); [Gopalan et al. \(2016\)](#); [Skrastins \(2017\)](#)). Within this literature, our paper is most directly related to research that studies the consequences of institutional variation across Native American reservations (e.g., [Karpoff and Rice \(1989\)](#); [Dippel \(2014\)](#)). The closest paper is [Brown et al. \(2017\)](#), which finds that stronger contract enforcement from PL280 led to an expansion of small business lending and bank branching on reservations, and that these effects translate into higher economic output in finance-dependent industries. [Brown et al. \(2017\)](#) also show that, unconditional on income, overall consumer credit scores are higher on PL280 reservations. The present paper shows that the variation in financial development across reservations documented by [Brown et al. \(2017\)](#) has important effects on consumer financial behavior. Notably, even after accounting for the higher incomes noted in [Brown et al. \(2017\)](#), we provide novel evidence that formative experiences with financial institutions lead to large differences in household credit management that persist well into adulthood. We also uncover unappreciated mechanisms through which these early life experiences affect consumer engagement with credit markets.

Finally, this paper adds to an important literature on the long-run effects of formative exposure to particular economic and institutional environments. For example, studying the financial behavior of immigrants to the United States, [Osili and Paulson \(2008\)](#) find that formative exposure to institutions that protect private property affects the propensity to own stock, while [Malmendier and Nagel \(2011\)](#) and [Malmendier and Nagel \(2016\)](#) show that formative experiences affect financial attitudes and beliefs about future macroeconomic conditions. [Malmendier et al. \(2011\)](#) show that formative experiences affect the personal decision-making of executives, and thus the effects of formative experiences extend to important corporate actions. In a distinct, but related vein, our work shows that a person's formative local financial environment has lasting effects on consumer credit behavior, and these effects persist for many years even after the person moves to a different institutional setting. Beyond the direct implications for our understanding of the importance of formative experiences, these findings are of particular interest given the importance of household debt and its effects on macroeconomic outcomes ([Favara and Imbs, 2015](#); [Baker, 2018](#)).

2. Reservation economies

2.1. Reservation institutions and Public Law 280

Native American reservations are an ideal setting to study how financial development affects household financial health. Reservations have a limited form of sovereignty in that they are generally not subject to state laws or regulations while still being subordinate to the rule of the US federal government. Arising from a federal policy commitment to tribal sovereignty, the historical status quo is

that each reservation runs its own tribal court to enforce the law on that reservation.⁶ In addition, reservations are relatively homogeneous on unmeasured dimensions due to similar long-term exposure to American institutions.

Although reservations have considerable political autonomy, the US Congress passed Public Law 280 in 1953, mandating that a subset of reservations in select states would be subject to jurisdiction by state courts.⁷ The reason not all reservations were assigned state jurisdiction under PL280 is that pre-existing disclaimers in many states' constitutions (established upon statehood) explicitly prohibit jurisdiction in reservation areas ([Anderson and Parker, 2017](#)). Thus, although court assignment under PL280 was not truly random, the ultimate geographic pattern of PL280 reservations is largely due to historical artifact.

Ultimately, PL280 was mandated in six states: California, Minnesota, Nebraska, Oregon, Wisconsin, and Alaska (upon statehood). In addition, Florida and Iowa asserted jurisdiction over their states' reservation court enforcement using a provision within PL280 that allowed states to opt into the law. Not all states that sought to opt into the law could effectively assert PL280 court jurisdiction. Indeed, between 1953 and 1968, eight other states asserted partial jurisdiction (e.g., applying only to pollution or highways), and other states attempted to assert PL280 authority but were constrained from doing so on account of provisions in their state constitutions ([Jimenez and Song, 1998](#)). Accordingly, even these optional cases were constrained by long-ago considerations at the inception of a state's constitution.⁸

According to legal scholars, PL280 was introduced because of a perceived need for stronger criminal enforcement on reservations, but state courts were also granted jurisdiction over civil contract enforcement "because it comported with the pro-assimilationist drift of federal policy and because it was convenient and cheap [to add to the law] ([Goldberg-Ambrose, 1997](#), p. 50)." Despite the intention to affect criminal enforcement, PL280 and non-PL280 areas had similar incarceration rates prior to the law's passage ([Brown et al., 2017](#)), and after the law, legal scholarship suggests the criminal provisions of PL280 were largely ineffective at reducing crime ([Dimitrova-Grajzl et al., 2014](#)). Thus, using PL280 assignment to evaluate the

⁶ A series of three Supreme Court cases decided by the Marshall Court, called the Marshall Trilogy (between the years 1823 and 1832), formalized this relationship between the US federal government, US states, and tribes. Congress has used the authority from the Marshall Trilogy to justify policy interventions on Native American reservations.

⁷ The law technically allowed for concurrent jurisdiction between state courts and tribal courts, but in effect, the introduction of state courts to reservations replaced tribal court activity on PL280 reservations.

⁸ Both Montana and North Dakota attempted to assert optional PL280 authority, but it did not come into force because it conflicted with their state constitutions. In separate legislation (Public Law 785 in 1950), New York reservations were subjected to the state court system. Because we want our measure to reflect whether state versus tribal courts have jurisdiction, we include New York reservations under our measure of state court jurisdiction but exclude reservations in Montana and North Dakota. In addition, several reservations were exempted from the original law, or had court authority retroceded to them, in which case we consider them under tribal court jurisdiction.

effects of financial development is not confounded by either preexisting or subsequent differences in law and order across reservations.

In all cases in which state courts were granted authority on reservations under PL280, the authority was granted to state courts without tribal consent. In 1968, Congress passed the Indian Civil Rights Act, which contained a provision requiring states to obtain tribal approval before any additional assertions of PL280 authority. Because tribes have been unwilling to relinquish sovereign control over their court systems, there have been no additional assertions of state court authority after the Indian Civil Rights Act.⁹ Consequently, PL280 caused persistent differences in reservation institutions that were not chosen by the tribes themselves.

To maintain the broadest possible sample for our empirical tests, we classify a reservation as under tribal courts if state courts cannot hear civil disputes on the reservation either because the reservation's state never asserted court jurisdiction over native lands or because PL280 jurisdiction was exempted or retroceded as outlined in the 1953 law or in the 1968 amendments to the law in the Indian Civil Rights Act. Otherwise, a reservation is considered to fall under state court jurisdiction. This definition is consistent with other studies on the consequences of PL280 (Anderson and Parker, 2008; Cookson, 2010; Parker, 2012).

2.2. Reservation financial development before and after Public Law 280

The historical narrative suggests that assignment to state courts under PL280 was unrelated to a reservation's financial or economic development, and recent studies show that initial conditions on reservations with state and tribal courts were not different in ways that could confound estimates of PL280's long-run impact. Specifically, Parker (2012) and Brown et al. (2017) show that credit markets, economic development, and demographics were broadly similar across state and tribal court jurisdictions prior to PL280's passage. We summarize this evidence in Panel A of Table A.1 in the online Internet Appendix. Notably, per capita incomes and unemployment rates are almost identical across state and tribal court reservation areas in the years immediately preceding PL280.

Most importantly for our study, local banking markets were also very similar across state and tribal court jurisdictions prior to the 1953 law. Brown et al. (2017, Table 1) use hand-collected data from the 1952 edition of Polk's Bank Directory (Polks) to construct county-level measures of banking activity (bank assets, bank loans, and total number of branches for banks headquartered in the county). According to the Polks data, per capita bank loans were not statistically different under state courts (\$201) versus tribal courts (\$192). Bank assets per resident were

also similar across jurisdiction (\$614 in state and \$597 in tribal court counties), as were the number of bank branches per capita in 1952 (0.248 per thousand under state versus 0.313 per thousand under tribal court counties). Parker's (2012) evidence on aggregate Bureau of Indian Affairs (BIA) regions also supports the conclusion that reservations targeted by PL280 had initial credit market conditions similar to tribal court reservations.

These similar initial conditions contrast with the stark differences in local financial development across reservations that emerged in the decades following PL280's passage (summarized in Table A.1, Panel B). Brown et al. (2017) find that several key indicators of subsequent banking market development are significantly greater on reservations with state courts, including the propensity for banks to extend credit and the extent of bank branching activity. For example, their estimates suggest state court jurisdiction led to approximately 20% more community bank branches per capita. On this basis, our specifications exploit these large differences in financial development, effectively using tribal court status (i.e., unaffected by PL280) to indicate low financial development.

Importantly, PL280 led to differences in financial development across reservations while not directly influencing most consumer financial products. PL280 primarily affects contracts related to secured local lending, not the unsecured debt used by most consumers to build a credit history (e.g., credit cards from national lenders). Mortgage lending is an important source of secured lending, but mortgage loans to reservations are fully guaranteed by the US Department of Housing and Urban Development, regardless of reservation court jurisdiction. Because PL280 does not directly complicate the legal provision of consumer credit, the variation in local credit market activity arising from PL280 allows our analysis to more credibly speak to the causal link between early exposure to local financial institutions and consumer financial health.

2.3. Event timeline and empirical strategy

Fig. 1 uses a timeline to illustrate our empirical strategy. In the approximately three decades following the 1953 passage of PL280, significant differences in local financial market development emerge across reservations. In the 1980s and 1990s, the individuals we study are born and grow up in different financial market environments. We then measure credit outcomes for these individuals over the 1999 to 2015 sample period, the time when they begin to build and manage credit histories as young adults.

The time lag between the 1953 enactment of PL280 and our sample period is crucial to our empirical design because we require sufficient long-run variation in financial development to estimate the exposure effects of growing up without finance. Most notably, the long-run differences in local financial development took several decades to fully emerge after the passage of the law. After the law change, it took time for subsequent court decisions to clarify the meaning of the law and for local financial institutions to respond to the law by expanding banking presence in these areas. Additionally, we seek to evaluate the effects of growing up in areas of relative financial underdevelop-

⁹ The 1968 Indian Civil Rights Act also allowed for retrocession of PL280 authority, but the process for retrocession of state court authority to tribal courts is difficult to initiate by tribes. Thus, there were few instances where tribal court authority was regained. We account for retrocession in our main measure as well as robustness to alternatives in related work (Brown et al., 2017).

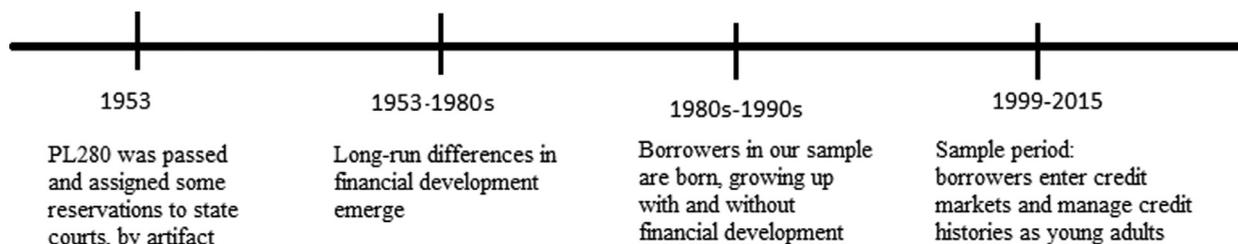


Fig. 1. Timeline of the empirical design.

This figure presents the timeline of events in our empirical design. It links the enactment of PL280 in 1953 to subsequent financial development on Native American reservations. These differences in financial development lead to cross-sectional differences in early life exposure to local financial institutions for young adults in our 1999–2015 sample of individuals in the Federal Reserve Bank of New York Consumer Credit Panel (FRBNY-CCP).

ment. As such, we not only need a lag between the legal change and the development of financial markets, but also time for individuals exposed to different levels of local finance to grow up and begin to enter formal credit markets. Thus, even if it were possible to gather micro-level consumer credit data around 1953, difference-in-differences tests around PL280's implementation would be uninformative about how consumer credit outcomes are affected by early life exposure to local financial development.

The central empirical challenge to this approach is to distinguish the effects of local financial development from broader changes to economic activity that also occurred following the passage of PL280 (Anderson and Parker, 2008; Brown et al., 2017). Our specifications take the following general form:

$$Y_{it} = \gamma_t + \gamma_{\text{census.region}} + \beta_1 \text{tribalcourt}_i + \beta_2 \text{income}_i + \epsilon_{it}, \quad (1)$$

where we observe household finance outcomes Y_{it} at the consumer i and quarter t level. As PL280 was passed at the state level, we employ census region fixed effects because these are the most granular geographic fixed effects that leave identifying variation in PL280 within region. Although census region fixed effects control for arbitrary regional differences in economic opportunity, it is possible there are important differences in economic activity within Census regions. We address this possibility by controlling for income in the census-tract where consumer i resides (income_i).

A potential shortcoming of Eq. (1) is that even income_i , measured at the census tract level, does not sufficiently control for differences in economic opportunity across tribal court and state court reservations. We confront this limitation by augmenting Eq. (1) with controls for indicators of human capital (i.e., high school and college education rates), other economic indicators (i.e., employment rates), socioeconomic factors (i.e., marriage rates), and other indicators of credit demand (i.e., credit scores of older generations not exposed to PL280 during formative years). These additional controls have minimal impact on our estimates of local financial development's effect on household financial health (β_1). We also perform a placebo test on reservation populations that were not exposed to differences in finance growing up but are presently subject to differences in economic circumstances—those born before PL280 came into effect. We find no difference in the financial health outcomes of older generations on reserva-

tions whose formative years were prior to PL280's 1953 passage. If there were uncontrolled-for differences in the level of long-run economic opportunities across reservations, we would expect individuals affected by these economic conditions to have different financial outcomes. We describe these tests in more detail in Section 5.1.

3. Data and measurement

3.1. Using census tract data to study reservation outcomes

To link reservations to household-level data, we compile a list of reservation area census tracts from the Tiger/Line American Indian/Alaska Native/Native Hawaiian Census geographic shape files. The FRBNY Consumer Credit Panel (FRBNY-CCP) reports the census tract location of sampled individuals at the time of the credit record. This provides a precise geographic mapping to consumers who reside in reservation areas. Thus, we are confident that our measures of consumer credit activity correspond to consumers who live on reservation lands and thus are directly exposed to the financial environments we discuss in Section 2. As Dimitrova-Grajzl et al. (2015) note when examining the FRBNY-CCP panel, this is an important advance in precision of data, given the data gaps in reservation areas.

Building on the sample of reservations (>250 residents in 1989) studied in Brown et al. (2017), our sample includes 367 reservation census tracts, 67 of which have state legal jurisdiction, and 300 fall under tribal courts. These census tracts are located on a total of 105 Native American reservations. Online Internet Appendix Fig. A.1 presents the location of US census tracts that have reservation lands. Reservations under PL280 status are noticeably scattered across regions of the United States. Online Internet Appendix Table A.2 presents the geographic distribution of consumers and observations in our sample (including reservation residents when they have left the reservation).

3.2. Data sources

3.2.1. Household financial activity

Our main data source is the FRBNY-CCP. This longitudinal data set tracks household liabilities and repayment using a 5% randomized sample of individuals with a social security number and a credit report on file at Equifax. The data start in 1999Q1 and are collected quarterly thereafter (our sample ends in 2015Q2). The sample design of

the Consumer Credit Panel alleviates concern over attrition: the panel resamples at every quarter to incorporate new credit report holders, and thus, is representative at any quarter. Further, as Brown et al. (2016) illustrate, the FRBNY-CCP offers a comprehensive coverage of US liabilities according to comparisons with other nationally representative surveys such as the the Flow of Funds Accounts and the Survey of Consumer Finances.

The FRBNY-CCP is particularly well-suited to studying household financial activity linked to reservations because of its scope of coverage (approximately one out of every 20 individuals who are 18 years or older is in the data) and the geographic precision assigned to the sampled consumers (census tract level). No other comprehensive data set on households (e.g., the Survey of Consumer Finances or the Panel Study of Income Dynamics) has the same geographic precision and coverage. The primary shortcoming of the FRBNY-CCP relative to other household surveys is that—aside from consumer age—there is no demographic information linked to the credit records, primarily due to federal laws prohibiting the use of race, sex, or national origin in the decision to extend credit. Although this restriction limits our ability to examine the heterogeneity in outcomes, owing to the data's random sampling and geographic precision, our main tests reliably estimate the effect of geographic exposure to financial development on the sample average of consumer credit outcomes.

3.2.2. Subsamples for the empirical analysis

To study the effect of financial development on consumer financial health, we focus on the subsample of FRBNY-CCP consumers whose first credit report corresponds to a residence on reservation lands, and we focus on young borrowers by retaining only records of individuals who are 18 years or younger at the start of the sample (1999). By focusing on borrowers who start their credit history on the reservation, our tests capture effects on consumer financial health precisely for the individuals who grew up on reservation land and thus were exposed to different levels of financial development.

In the empirical analysis, we seek to understand how financial development influences early life financial outcomes, the importance of these effects in the long term, and how persistent these effects are when an individual moves from an underdeveloped area. To this end, our empirical tests focus on the following subsamples: (1) the sample of young borrowers (aged ≤ 25), which enables us to focus directly on early life entrance into financial markets; (2) the sample of relatively older borrowers (aged ≥ 25) who remain on the reservation for the entire sample period, which allows us to estimate the long-run effects of financial development on financial health; and (3) the sample of borrowers who moved from the reservation to an off-reservation location, which enables us to evaluate the speed of recovery from moving to an area with stronger financial development. We join other papers, such as Dettling and Hsu (2014) and Dokko et al. (2015), that exploit the longitudinal features of the FRBNY-CCP by providing separate tests for consumers who stay on reservations for the entire sample and by studying the dynamics of consumer financial health for those who

move away. Aside from speaking to dynamics, studying differences between consumers who stay on reservations and those who move away helps isolate the impact of different institutional settings experienced during one's formative years on subsequent outcomes.

3.2.3. Outcome variables from the credit bureau data

The analysis focuses on several key variables from the FRBNY-CCP, for which summary statistics are presented in Table 1, Panel A. Our primary measure of consumer financial health is the Equifax riskscore, which varies between 280 and 850 and is similar to a consumer's FICO score. We focus on the riskscore because it is a nationally standardized measure that summarizes an individual's history of borrowing and repayment activity. As a direct measure of consumer creditworthiness, lenders use metrics like the Equifax riskscore in the decision to extend credit, as well as the interest rates they charge. Thus, a higher riskscore can lead to significant cost savings on loans and increased consumer welfare.

A useful alternative measure of financial health is the presence of delinquent accounts, which we measure by calculating the fraction of credit accounts (tradelines) that are at least 90 days past due. Specifically, the variable equals the number of credit accounts 90 days past due, 120 days past due or in collections, or severe derogatory, divided by the total number of credit accounts in the current quarter. A key advantage of this measure is that it captures how well borrowers manage their credit, conditional on having obtained credit in the first place. In this way, evidence on delinquencies speaks to the concern that differences in the Equifax riskscore are driven entirely by unaccounted-for differences across reservations in the opportunity to obtain credit. That said, focusing only on delinquencies overlooks some important aspects of consumer financial health because, as we find in the following section, PL280 has significant effects on consumer entry into formal credit markets.

3.2.4. Individual survey

We complement our tests that use the FRBNY-CCP by commissioning Qualtrics to survey individuals on personal attributes and financial attitudes and aptitude. To parallel the FRBNY-CCP sample, we instructed Qualtrics to collect a sample of respondents that identify as Native American, say that they grew up near reservation lands, and who are 40 years old or younger. Our main sample includes 704 responses after manually removing invalid responses that indicate that the respondent is not Native American (e.g., reporting nonsensical tribal affiliations) and responses that were completed too quickly or slowly relative to the expected duration of the survey. For comparison to the broader population, we collected a comparison sample of 125 respondents (also 40 years old or younger), evenly split between men and women, that matches the racial composition of the entire United States.

Our survey includes questions on financial attitudes and aptitude that are commonly used in the literature (Table 1, Panel B presents summary statistics, and the survey in the online Internet Appendix presents the exact language of the survey questions). Specifically, we

Table 1

Summary statistics for regression analysis.

Panel A presents summary statistics of data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax.

The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for consumers who were 18 or younger in 1999 and whose first credit report was on reservation lands as defined by the Bureau of Indian Affairs. The variable *tribalcourt* equals one if the consumer resides on a reservation using tribal courts as determined by Public Law 280. The variable *offresvn* equals one if the observation comes from a quarter in which the consumer resides off reservation lands. Each observation is at the consumer-quarter level. Panel B presents data from the Consumer Financial Protection Bureau's database of consumer complaints (December 2011 to July 2017). We match complaints in the data to ZIP codes on Native American reservations. Each observation is a single complaint. Panel C presents data from a novel survey administered by Qualtrics. The survey was targeted to respondents 40 years old or younger who identify as Native American and grew up on or near a Native American reservation. Responses were collected between January 2018 and June 2018. Each observation corresponds to one individual respondent.

Panel A: Credit bureau data (FRBNY-CCP)						
	observations	mean	median	std dev	10th percentile	90th percentile
<i>Sample: all consumer - quarter observations, on and moved from reservation</i>						
Equifax riskscore	350,798	635.7	645	93.1	512	754
# credit inquiries during past 12mo	278,047	2.35	2	2.66	0	5
fraction delinquent (> 90 days past due)	285,925	0.14	0	0.33	0	1
tribalcourt (= 1)	350,798	0.77				
off resvn (= 1)	350,798	0.51				
<i>Sample: 25+ years old, on reservation entire sample</i>						
Equifax riskscore	45,320	624.1	615	90.8	615	750
# credit inquiries during past 12mo	32,381	1.91	1	2.32	0	5
fraction delinquent (> 90 days past due)	31,795	0.24	0	0.39	0	1
tribalcourt (= 1)	45,320	0.82				
Panel B: Consumer complaint database (CFPB)						
	observations	mean	median	std dev	10th percentile	90th percentile
credit information issue (= 1)	1604	0.123				
no relief on complaint credit information issue (= 1)	197	0.726				
tribalcourt (= 1)	1,604	0.668				
Panel C: Qualtrics custom survey						
	observations	mean	median	std dev	10th percentile	90th percentile
age obtained first card	268	21.84	20	4.929	18	29
financial literacy	704	1.435	1	0.931	0	3
stock/bond literacy	704	0.464	0	0.609	0	1
financial trust	704	3.268	3	1.769	1	5
general trust	661	0.121	0	0.326	0	1
age	704	28.03	28	6.164	20	37
has a credit card (= 1)	704	0.381				
grew up near bank (= 1)	704	0.666				
male (= 1)	704	0.325				
level of education (N = 704)	some high school	high school graduate	tech. or prof. program	some college	college graduate	
	29.40%	57.24%	2.27%	7.39%	3.69%	
level of income (N = 704)	0 to \$10K	\$10K to \$20K	\$20K to \$30K	\$30K to \$50K	\$50K to \$100K	greater than \$100K
	37.64%	21.31%	15.77%	15.34%	8.10%	1.85%

use measures of financial literacy drawn from [Lusardi and Mitchell \(2011\)](#) and measures of trust taken from [Guiso et al. \(2008\)](#). We measure risk tolerance similar to other household surveys, such as the Survey of Consumer Finances, and follow [Dimmock et al. \(2016\)](#) to measure ambiguity aversion. The survey also allows for us to measure income and education levels at the individual level, which is not possible in the FRBNY-CCP. We use our nationally representative comparison sample to validate the quality of our survey responses. In particular, evidence from the Financial Industry Regulatory Authority (FINRA) Investor Education Foundation's 2015 National Financial Capability Study (FINRA 2017) shows that Native Americans have less financial knowledge than the broader population. This is indeed what we find ([Fig. 8](#)).

We use the survey to complement our main tests by providing evidence regarding *why* formative exposure to banking matters for consumer credit profiles. Thus, our

tests relate financial attitudes and aptitudes to how consumers respond to the question, "Was there a bank in the community where you grew up?". This question provides a direct measure of exposure to banking, which we relate to financial attitudes and aptitudes along different dimensions.

4. Engagement with consumer credit markets

4.1. Empty credit records

[Fig. 2](#) provides evidence of the significant differences in consumer engagement with credit markets across state and tribal court jurisdictions. For each census tract in the sample, we calculate the number of FRBNY-CCP credit reports for consumers younger than 25 in the quarter divided by the tract's population 25 years or younger according to the 2000 Census. Because the FRBNY-CCP is a 5% random sample, we multiply this ratio by 20 to get an es-

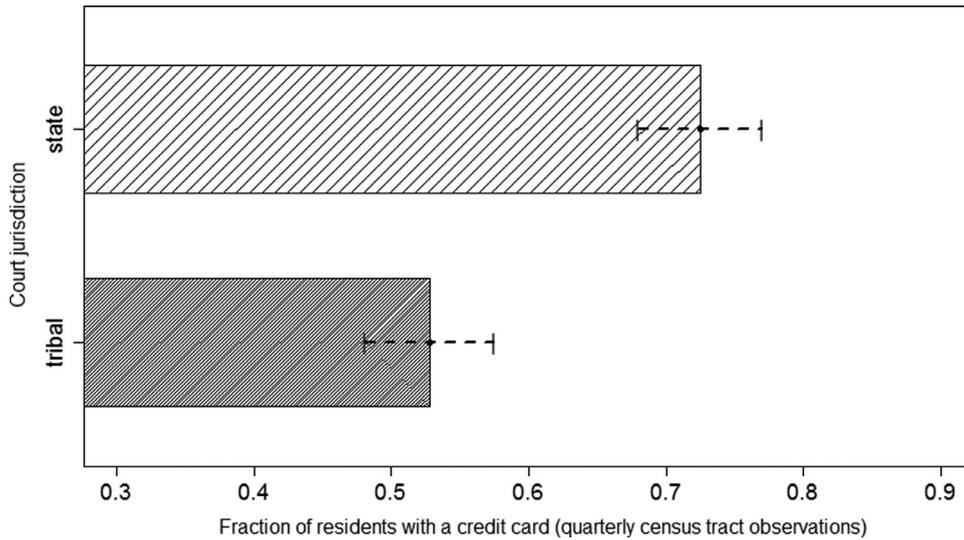


Fig. 2. Credit undercoverage across reservations.

This figure demonstrates the extent of credit undercoverage on reservations. For each census tract (quarterly) in the sample, we calculate the number of FRBNY-CCP credit reports for consumers 25 years or younger divided by the tract’s population 25 years or younger according to the 2000 census. Because the FRBNY-CCP is a 5% random sample, we multiply this ratio by 20 to get an estimate of the proportion of individuals with a credit report. The figure presents the median census tract (and 95% confidence interval for the median) on state jurisdiction reservations (civil contracts are adjudicated in state’s court system) or tribal court reservations (civil contracts are adjudicated in reservation tribal courts).

estimate of the proportion of individuals with a credit report. Fig. 2 shows that for the median state-court-reservation census tract, approximately 73% of individuals have a credit report, whereas the corresponding value for tribal court reservations is only 53%. This gap in credit coverage is statistically significant (dashed lines represent the 95% confidence intervals).

4.2. Consumer age when entering credit markets

4.2.1. Graphical evidence

Fig. 3 shows that it takes longer for individuals to formally enter credit markets if they grow up in areas falling under tribal court jurisdiction. The figure plots the proportion of the sample receiving their first credit account (line of credit) at any given age. A smaller fraction of 18- and 19-year-old individuals receive their first line of credit under tribal courts. Approximately 47% of consumers who eventually receive a credit account do so by age 19 under state courts versus 39% of consumers who receive a credit account under tribal courts. As a consequence, a larger share of the tribal court consumers who eventually obtain credit receive their first line after age 20. There are also differences between young borrowers in tribal court versus state court reservations in the age at which consumers receive their first credit report. The main difference between a first report and a first line of credit is that a consumer can have a credit report even if they do not yet have any (approved) credit accounts.

4.2.2. Regression estimates

To study the effect of financial development on the propensity to enter credit markets, we estimate the follow-

ing linear probability model:

$$first\ credit_{it} = \gamma_t + \gamma_{census\ region} + \gamma_{age} + \beta_1 tribal\ court_i + \beta_3 income_i + \epsilon_{it} \quad (2)$$

The dependent variable *first credit* equals one in the first quarter that consumer *i* receives their first credit report (or line of credit) and is equal to zero in all preceding quarters starting with the year in which *i* turned 18 years old. The specification includes fixed effects for *i*’s age in quarter *t* and fixed effects for quarter *t*. Hence, the specification is set up like a Cox proportional-hazards model, which we instead estimate using Ordinary Least Squares.¹⁰ The independent variable of interest, *tribalcourt*, equals one if the consumer resides on a reservation using tribal courts as determined by PL280, and the variable *income* is the year 2000 median household income for the consumer’s census tract.

Table 2 presents estimates of Eq. (2) using the FRBNY-CCP credit records for consumers whose first credit report is on reservation lands. Columns 1a and 2a in Panel A report regression specifications for the time until *i*’s first line of credit. The estimate of β_1 in column 1a is -0.60% (statistically significant at the 1% level). Accordingly, *tribalcourt* lowers the probability of receiving a first line of credit by 13% (0.6% divided by the unconditional probability of *first credit* – 4.6%). The coefficient estimate is similar (-0.5%) after including indicators for the nine census sampling regions (column 2a). Hence, *tribalcourt* has a negative effect on the propensity to enter credit markets.

¹⁰ We use linear probability models instead of Cox proportional-hazards because in subsequent tests we interact *tribalcourt* with other variables. Interaction terms estimated in nonlinear models are difficult to interpret.

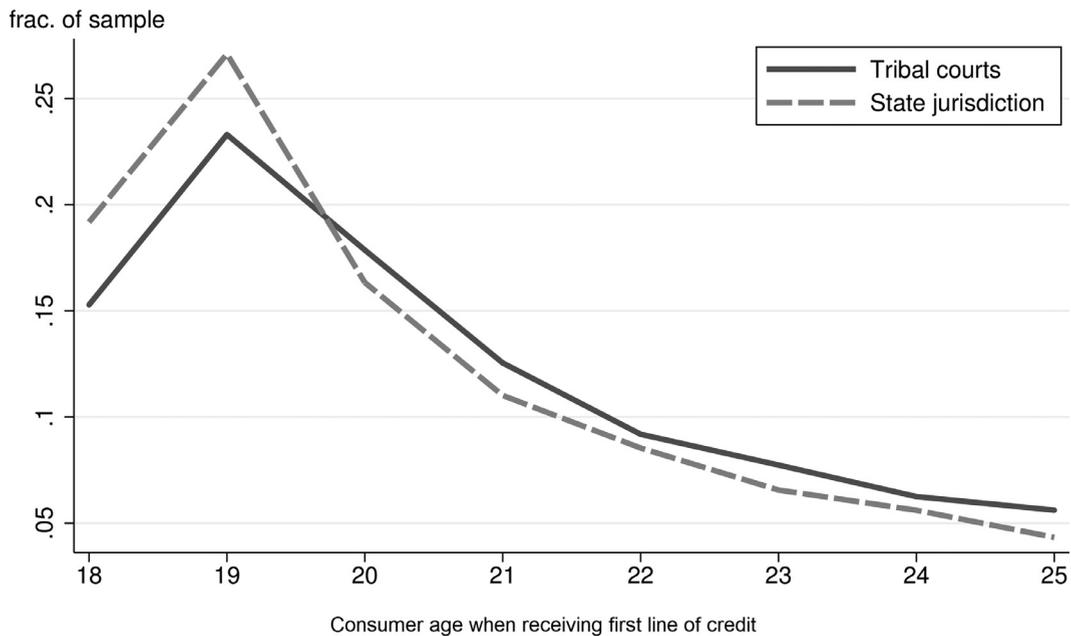


Fig. 3. Delayed access to credit.

This figure uses data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for borrowers who would have turned 18 by 1999 and whose first credit report is associated with an address on Native American reservation lands as defined by the Bureau of Indian Affairs. In state jurisdiction reservations, civil contracts are adjudicated in the state's court system, as prescribed according to Congressional legislation titled Public Law 280. In tribal court reservations, the tribe's court system adjudicates and enforces civil contracts. *Consumer age when receive first line of credit* equals the consumer's age when they obtain their first line of credit.

For robustness, columns 3a and 4a show that residents of tribal court reservations also take longer to obtain a first credit report. Using Eq. (2) to estimate the hazard to the consumer's first credit report, the estimate of β_1 equals -0.93% ($s.e. = 0.0018$). The estimated relationship between time to obtain a first credit report and *tribalcourt* is also negative after including the census region indicators, though the magnitude and statistical significance are weaker. The coefficient estimate is -0.38% with census region effects, which is statistically significant at the 10% level (column 4a). Overall, these estimates show the likelihood that an individual in tribal court areas establishes a first credit report at age t is around 4%–10% lower than a corresponding individual located on a reservation with state courts (the unconditional probability of receiving a first credit report in columns 3a and 4a is 9.5%).

4.2.3. Evidence on the role of exposure to bank branches

We provide additional evidence that these differences in inclusion in formal credit markets are caused by differences in local financial development and not another omitted factor. In particular, any exogenous factor that increases the supply of local banking would cut against the differences between tribal and state court institutions. Near the beginning of our sample period, states were gradually relaxing regulations against interstate bank branching in accordance with the IBBEA. IBBEA led to large increases in bank branch density, particularly in low-income and rural areas, but had little effect on economic growth (Celerier and Matray, 2016). Thus, we expect the increase in bank

branch density associated with IBBEA to partially offset tribal court's effect on local financial development.

According to regression estimates in Panel B of Table 2, the lower financial inclusion on reservations with tribal courts is mitigated by bank branching expansion after IBBEA. These specifications estimate the same linear probability model as in Eq. (2) separately for the subsample of borrowers from reservations in IBBEA-deregulated states and for states that did not deregulate banking and branching under IBBEA prior to the current quarter.¹¹ We define a state as being deregulated if it adopted any of the four pro-deregulation bank branching policies described by Rice and Strahan (2010).

For both subsamples, the coefficient estimate on *tribalcourt* is negative and statistically significant at the 5% level or better. Moreover, comparing the results from the non-deregulation (column 1b) and deregulation (column 2b) samples, the deregulation-induced expansion of bank branch supply in the surrounding state partially mitigates the influence of local financial development on the time to a consumer's first approved line of credit. Specifically, the estimated effect of *tribalcourt* on time to first account in the deregulated sample is approximately 50% smaller than the magnitude of the effect on the non-deregulated sample. Column 3b includes an interaction between *tribalcourt* and an indicator for the state being deregulated. This specification confirms our split sample results—the in-

¹¹ Because IBBEA status changes in a small number of states during the sample period, there are a few consumers that are part of both split samples.

Table 2

How long does it take to enter credit markets? Linear probability estimates. This table presents estimation results from the linear probability model

$$first\ credit_{it} = \gamma_t + \gamma_{census\ region} + \gamma_{age} + \beta_1 tribalcourt_i + \beta_2 income_i + \epsilon_{it}$$

using data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes credit records between 1999Q1 and 2015Q2 for borrowers whose first credit report is associated with a census tract on reservation lands as defined by the Bureau of Indian Affairs. The dependent variable *first credit* equals one in the first quarter that consumer *i* receives their first line of credit or first credit report (the regressions exclude observations that occur after *i* has received their first line of credit or first credit report). *Tribalcourt* equals one if the consumer's first credit report is on a reservation using tribal courts as determined by Public Law 280. *Median income* comes from the 2000 US Census and is at the census tract level. Panel A presents the estimates for the full sample. In Panel B, the data are sorted by the state's time-varying status of deregulation under the Interstate Banking and Branching Efficiency Act of 1994. We call the state deregulated if the state has adopted as of quarter *t* any of the four bank branching measures described in Rice and Strahan (2010). Standard errors are clustered by census tract and date. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

Panel A: Exposure to banks and credit market entry				
<i>t</i> = age – 18	time to first line of credit		time to first credit report	
	(1a)	(2a)	(3a)	(4a)
<i>tribalcourt</i>	–0.00601*** (0.0012)	–0.00510*** (0.0014)	–0.00925*** (0.0018)	–0.00378* (0.0022)
<i>median income / 1000</i>	0.000651*** (0.000068)	0.000644*** (0.000067)	0.000688*** (0.000082)	0.000697*** (0.000082)
date-quarter FE	x	x	x	x
age FE	x	x	x	x
census region FE		x		x
mean of dep var	0.046	0.046	0.095	0.095
<i>N</i> (consumer-quarter)	246,735	246,735	151,394	151,394
<i>N</i> (consumers)	14,380	14,380	14,380	14,380
<i>R</i> ²	0.047	0.047	0.12	0.12

Panel B: The role of bank branching expansion							
<i>t</i> = age – 18	IBBEA status:	time to first line of credit			time to first credit report		
		not deregulated	deregulated	all states	not deregulated	deregulated	all states
		(1b)	(2b)	(3b)	(4b)	(5b)	(6b)
<i>tribalcourt</i>		–0.0107** (0.0052)	–0.00551*** (0.0014)	–0.0386* (0.021)	–0.0273*** (0.010)	–0.00341 (0.0023)	–0.108** (0.042)
<i>tribalcourt</i> × <i>deregulated</i>				0.0400* (0.021)			0.0976** (0.042)
<i>median income / 1000</i>		x	x	x	x	x	x
date-quarter FE		x	x	x	x	x	x
age FE		x	x	x	x	x	x
census region FE			x			x	
state FE				x			x
mean of dep var		0.042	0.046	0.046	0.088	0.096	0.095
<i>N</i> (consumer-quarter)		25,197	221,538	246,735	16,404	134,990	151,394
<i>N</i> (consumers)		1,471	12,958	14,380	1,471	12,935	14,380
<i>R</i> ²		0.039	0.049	0.048	0.14	0.12	0.12

teraction term is positive and statistically significant. The specification is useful because we can include state fixed effects, which help control for any time-invariant unobservable features of the states that chose to or chose not to deregulate their banking sector. These differences are even more pronounced if we estimate the linear probability model using time to first credit report as the dependent variable (columns 4b, 5b, and 6b). Overall, these results support the inference that greater access to local financial institutions helps consumers more quickly integrate into formal credit markets.

4.3. Exposure to banks and credit card usage

Our Qualtrics survey provides complementary evidence that exposure to banks when growing up increases engagement with credit markets. Table 3 presents estimates of

the following OLS regression:

$$creditcard_i = \beta_1 grow\ up\ near\ bank_i + \beta_2 controls_i + \epsilon_i. \quad (3)$$

The dependent variable *creditcard_i* in columns 1 and 2 is an indicator variable for whether the respondent currently or has ever had a credit card. In columns 3 and 4, the dependent variable *creditcard_i* is the respondent's age upon receiving his or her first credit card. The independent variable of interest, *grow up near bank*, equals one if the respondent answers yes to the question, "Was there a bank in the community where you grew up?" All specifications include controls for the respondent's gender and age as well as census region fixed effects.

The regression estimates show that individuals growing up near a bank are between seven to ten percentage points more likely to have a credit card. This is a large effect relative to the fraction of individuals in the Qualtrics survey who have a credit card (38%). We also find that,

Table 3

Exposure to banks and credit card usage, survey evidence.

This table uses data from a survey administered by Qualtrics. The survey was targeted to respondents 40 years old or younger who identify as Native American and grew up on or near a Native American reservation. This table presents OLS estimation results of the following specification

$$\text{creditcard}_i = \beta_1 \text{grow up near bank}_i + \beta_2 \text{controls}_i + \varepsilon_i.$$

The dependent variable in columns 1 and 2 is equal to one if respondents answer yes to the question, "Have you ever had or presently have a credit card?" In columns 3 and 4, the dependent variable equals the age when the respondent first received a credit card. The independent variable of interest, *grow up near bank*, equals one if the respondents answers yes to the question, "Was there a bank in the community where you grew up?" Standard errors are clustered by state. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

dep var:	has a credit card (= 1)		age first credit card	
	(1)	(2)	(3)	(4)
grow up near bank (= 1)	0.110*** (0.034)	0.0727** (0.033)	-1.262** (0.52)	-1.346** (0.53)
male	0.0116 (0.036)	0.00933 (0.034)	1.215** (0.49)	1.355** (0.58)
age	0.0251 (0.028)	0.0268 (0.023)	1.142** (0.51)	1.271** (0.51)
age squared	-0.000339 (0.00051)	-0.000373 (0.00043)	-0.0138 (0.0091)	-0.0161* (0.0091)
Census Region FE	x	x	x	x
Income FE		x		x
Education FE		x		x
Number of respondents	704	704	268	268
R ²	0.031	0.096	0.22	0.25

conditional on having a card, respondents growing up near a bank are 1.3 years younger upon receiving their first card. These results are robust to controlling for income and education at the individual (respondent) level. This evidence supports our inferences from the FRBNY-CCP data, and shows that the positive relation between formative exposure to banks and engagement with credit markets is not an artifact of our use of PL280 status to measure local financial development.

5. Formative exposure to finance and household financial health

This section examines the financial health consequences of formative exposure to finance. Unless otherwise indicated, the tests in this section focus on the sub-sample of borrowers who are at least 25 years old in quarter t and who we only observe residing on the reservation during the FRBNY-CCP sample. By focusing on the over-25 sample, our tests allow enough of the life-cycle to pass for the effects of delayed exposure to accumulate, and we measure financial health approximately when consumer finances start to significantly matter to individuals (e.g., when some households consider home ownership and family formation).

We find that more robust local financial markets are positively associated with better financial health. The top panel of Fig. 4 plots the tract-by-quarter distribution of credit scores by reservation type across our entire sample. Notably, the mass of high riskscores is larger for state court reservations than for tribal court reservations. Likewise, there is a larger fraction of subprime borrowers (riskscore <620) under tribal courts. Similarly, the bottom panel presents a bar graph, which shows that the delinquency rate is, on average, eight percentage points higher across tribal court census tracts.

5.1. Financial health of borrowers across reservations

We confirm the effects of local financial development on financial health using a regression analysis of consumer riskscores and delinquencies. Table 4 presents estimates of the following empirical specification,

$$Y_{it} = \gamma_t + \gamma_{\text{census region}} + \gamma_{\text{birthyear}} + \gamma_{\text{age}} + \beta_1 \text{tribalcourt}_i + \beta_2 \text{income}_i + \varepsilon_{it}. \quad (4)$$

The outcome variable Y_{it} is consumer i 's riskscore in Panel A and the fraction of delinquent accounts in Panel B. In addition to controlling for *income*, we subject the relation between consumer financial health and *tribalcourt* to an increasingly rich set of geographic and time-varying fixed effects. Column 1 includes quarter, birth year, and age fixed effects, while column 2 includes census region fixed effects, and column 3 interacts the census region and quarter fixed effects to allow for time variation in the effects of regional economic activity on consumer credit outcomes.

In Panel A, the coefficient estimate on *tribalcourt* is between -7 and -10 riskscore points and is statistically significant at the 1% level across all specifications. Notably, this effect of *tribalcourt* is over and above the effect of *income*, which, not surprisingly, has a strong positive impact on consumer riskscore. Comparing the coefficient estimates on *tribalcourt* and *income*, the independent effect of growing up on a tribal court reservation on credit scores is similar to the effect of reducing annual personal incomes by nearly \$6000.¹²

¹² The coefficient on *income*, β_2 , equals approximately 1.2, which means that every \$1000 increase in census tract median incomes increases average consumer riskscores by 1.2 points. Taking the lower end of estimates on *tribalcourt* (β_1 approximately equal to 7), means that median incomes would have to increase by \$5800 (7 divided by 1.2 times \$1000) to equal the effect of *tribalcourt* on riskscores.

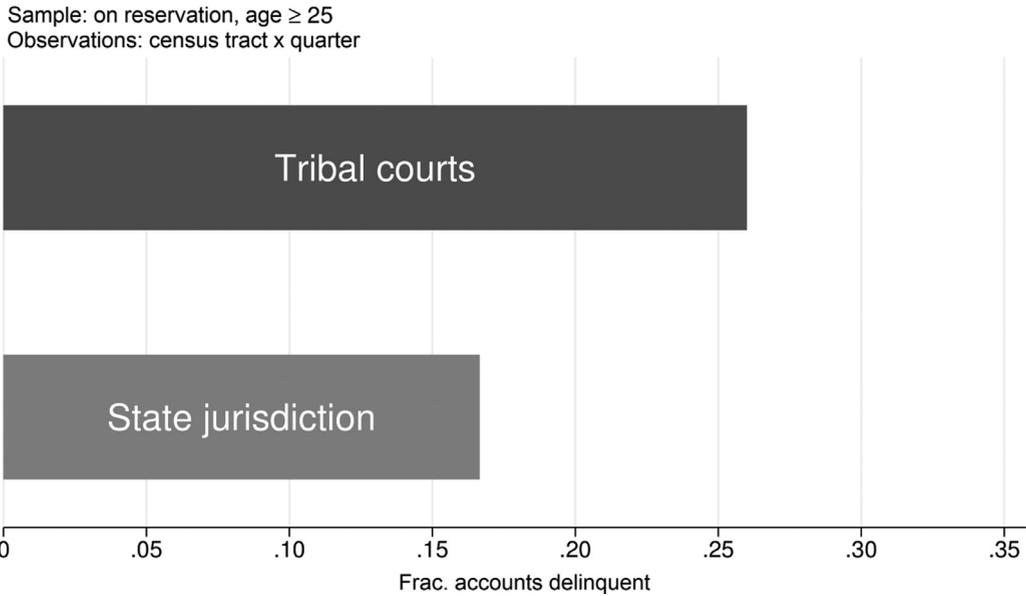
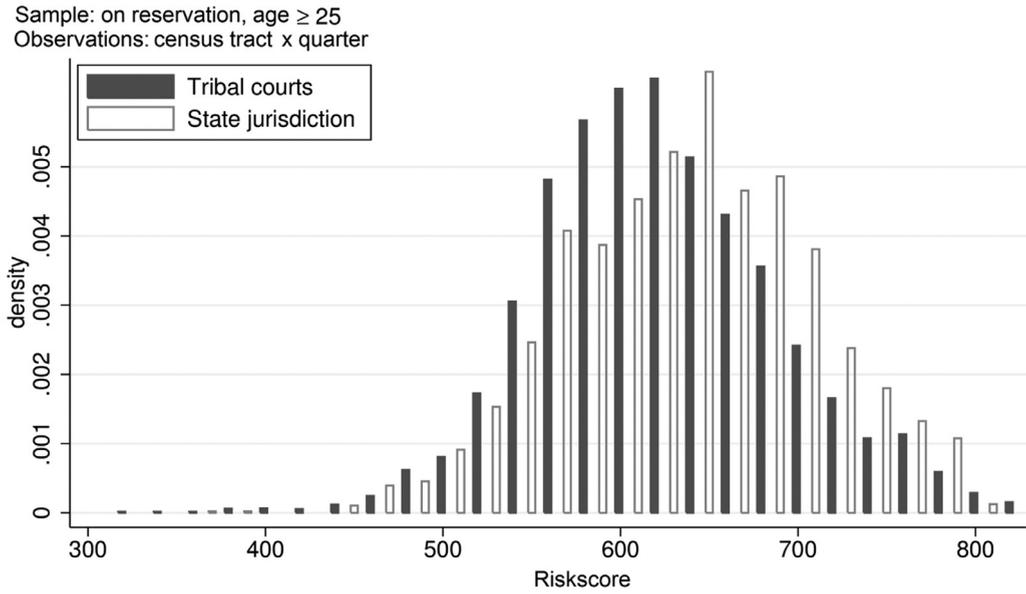


Fig. 4. The financial health of reservation borrowers in adulthood—graphical evidence. This figure uses data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for borrowers who were 18 years old or younger in 1999, are at least 25 years old in date t , and who reside only on reservation lands during the FRBNY-CCP sample. State court reservations are reservations for which civil contracts are adjudicated in the state’s court system, as prescribed according to Congressional legislation titled Public Law 280. In tribal court reservations, the tribe’s court system adjudicates and enforces civil contracts. *Riskscore* is a proprietary metric from Equifax that measures a consumer’s creditworthiness. It ranges from 280 to 850. *Frac. delinquent accounts* is the number of accounts at least 90 days past due divided by the total number of credit accounts on the consumer’s credit report.

To provide further context for the magnitude of the *tribalcourt* effect, the estimated effect of seven to ten riskscore points is large in comparison to individual-specific factors that affect consumer financial health. Referring to estimates in the literature, legal interventions

into consumer balance sheets, such as enforcing Chapter 13 bankruptcy protections, improve average credit scores by 14.9 points (Dobbie et al., 2017). The effect of local financial development on consumer financial health is in the neighborhood of the effect of Chapter 13 protections,

Table 4

The financial health of reservation borrowers upon reaching adulthood, consumers 25 years old and older. This table presents OLS estimation results of the following specification,

$$Y_{it} = \gamma_t + \gamma_{\text{census region}} + \gamma_{\text{birthyear}} + \gamma_{\text{age}} + \beta_1 \text{tribalcourt}_i + \beta_2 \text{income}_i + \epsilon_{it},$$

using data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for borrowers who turn 18 after 1999, and who are 25 or older as of quarter t . The sample contains consumers whose first credit report has an address on Native American reservation lands and who only appear on reservation lands during the sample. In Panel A, *riskscore* is a proprietary metric from Equifax that measures a consumer's creditworthiness. It ranges from 280 to 850. In Panel B, the dependent variable *frac accounts delinquent* equals the number of accounts at least 90 days past due divided by the total number of credit accounts. *Tribalcourt* equals one if the consumer's first credit report is on a reservation using tribal courts as determined by Public Law 280. *Median income* comes from the 2000 US census and is at the census tract level. Standard errors are clustered by census tract and date. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

Panel A: Financial health measured by <i>riskscore</i>			
	(1a)	(2a)	(3a)
<i>tribalcourt</i>	−10.14*** (1.28)	−7.521*** (1.41)	−7.035*** (1.42)
<i>median income / 1000</i>	1.296*** (0.035)	1.219*** (0.041)	1.201*** (0.041)
birth year FE	x	x	x
age FE	x	x	x
date quarter FE	x	x	
census region FE		x	
census region × date quarter FE			x
<i>N</i>	45,320	45,320	45,284
<i>R</i> ²	0.070	0.082	0.090
Panel B: Financial health measured by <i>frac. accounts delinquent</i>			
	(1b)	(2b)	(3b)
<i>tribalcourt</i>	0.0413*** (0.0061)	0.0207*** (0.0069)	0.0187*** (0.0069)
<i>median income / 1000</i>	−0.00441*** (0.00018)	−0.00408*** (0.00019)	−0.00406*** (0.00019)
birth year FE	x	x	x
age FE	x	x	x
date quarter FE	x	x	
census region FE		x	
census region × date quarter FE			x
<i>N</i>	31,795	31,795	31,795
<i>R</i> ²	0.042	0.050	0.060

which speaks to the strength of the behavioral mechanism we identify.

We also quantify the magnitude of the *tribalcourt* estimate by calculating how its effect on *riskscore* changes the monetary cost of a typical financial product. Specifically, we consider the costs to a 30-year-old borrower who seeks a \$100,000 30-year mortgage. To account for heterogeneous effects on consumer riskscores and discrete jumps in the cost of financing for different risk tranches, online Internet Appendix A.2 presents quantile regression estimates of changes at different points in the distribution of consumer riskscores. We estimate that local financial development shifts the consumer riskscore distribution such that six percentage point fewer borrowers are eligible for a conventional-rate mortgage on tribal court reservations. For borrowers who are eligible for a mortgage (half of the sample), the effect of local financial development shifts the riskscore distribution such that the cost of financing is 5% higher for tribal court residents than for state court residents.

Panel B shows that borrowers growing up on tribal court reservations have between two and four percentage points more delinquent accounts than similar borrow-

ers on state court reservations, even after controlling for median income in the borrower's census tract. These tests capture a different dimension of financial health than overall *riskscore* and are especially valuable for distinguishing household credit management from factors that influence credit demand because they condition on the consumer having already demanded and obtained credit.¹³

To address the potential concern that controlling for *income* does not adequately account for omitted factors affecting the demand for credit, we examine the robustness of the findings to including proxies for a range of alternative omitted factors (Fig. 5). In each case, whether we include controls for the area's economic characteristics

¹³ A reservation's level of banking activity has a similar relationship with financial health as the *tribalcourt* dummy variable. Online Internet Appendix Table A.4 shows that after controlling for *income*, a standard deviation increase in per capita number of bank branches in the county (Summary of Deposits, FDIC) is associated with an increase of approximately nine points in consumer riskscores. Moreover, a higher density of bank branches is associated with a significantly lower frequency of delinquent accounts. These findings are consistent with our estimates of tribal court's strong effect on financial health reflecting the workings of local financial development.

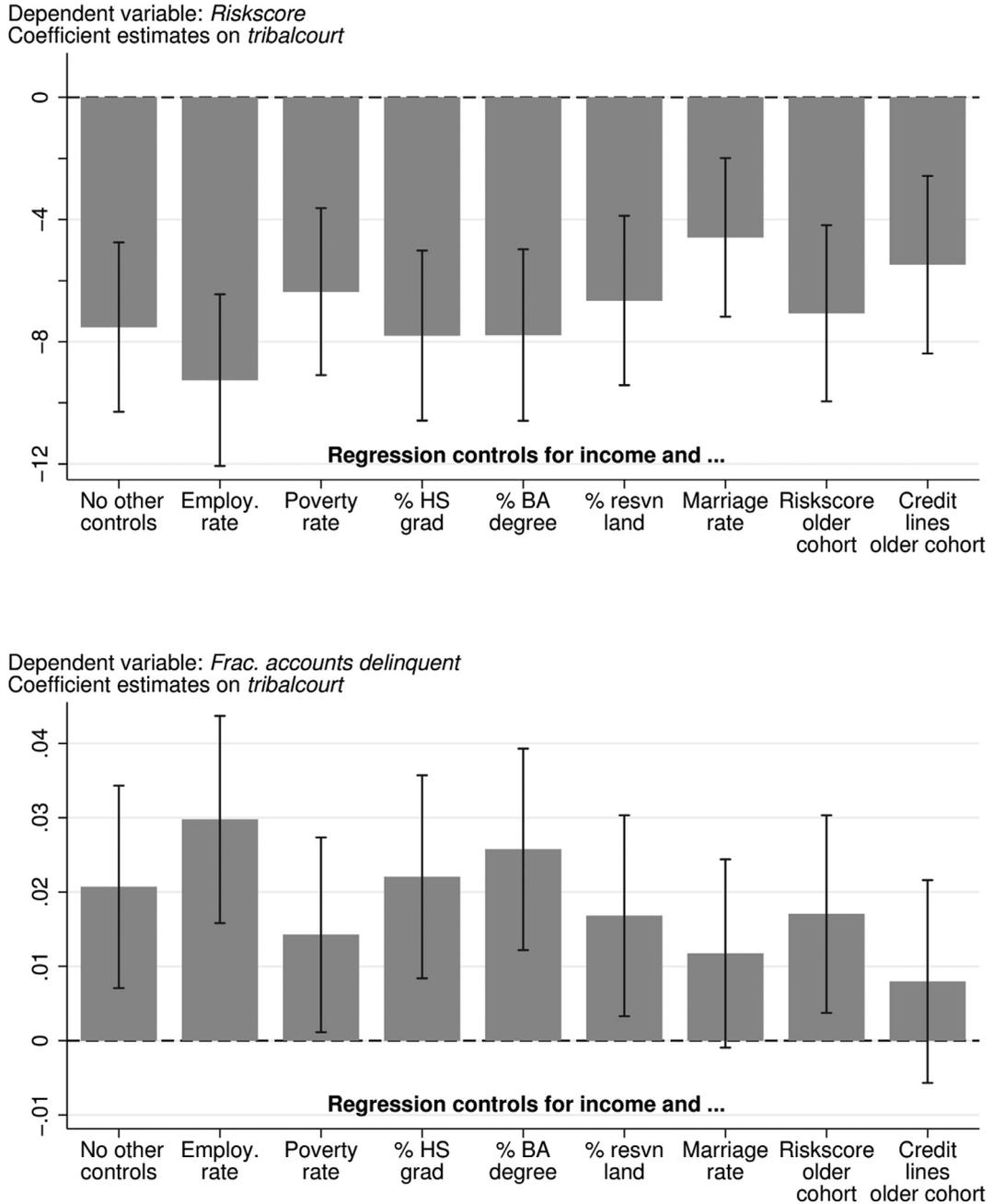


Fig. 5. The financial health of borrowers upon reaching adulthood—robustness. This figure is an extension of Table 4. It presents OLS estimation results of the following specification,

$$Y_{it} = \gamma_t + \gamma_{census\ region} + \gamma_{birth\ year} + \gamma_{age} + \beta_1 tribalcourt_i + \beta_2 income_{it} + \beta_3 control\ var_i + \epsilon_{it}$$

for consumers 25 years or older in period t . In the top figure, the dependent variable, *riskscore* is a proprietary metric from Equifax that measures a consumer’s creditworthiness. It ranges from 280 to 850. In the bottom figure, the dependent variable *frac accounts delinquent* equals the number of accounts at least 90 days past due divided by the total number of credit accounts. The variable *tribalcourt* equals one if the consumer’s first credit report is on a reservation using tribal courts as determined by Public Law 280. *Median income, employment rate, poverty rate, high school (HS) graduation rate, fraction of residents with a bachelor’s (BA) degree*, and the *marriage rate* come from the 2000 US Census and is at the census tract level. The *% reservation lands* comes from merging census tracts with Tiger/Line American Indian/Alaska Native/Native Hawaiian Census geographic shape files. *Riskscore* and *credit lines of older cohort* are the census tract averages of reservation consumers born between 1930 and 1953. Standard errors are clustered by census tract and date, and the figure presents 95% confidence intervals.

Table 5

The financial health of pre-PL280 birth year consumers.

This table presents OLS estimation results of the following specification,

$$Y_{it} = \gamma_t + \gamma_{\text{census region}} + \gamma_{\text{birthyear}} + \gamma_{\text{age}} + \beta_1 \text{tribalcourt}_i + \beta_2 \text{income}_i + \epsilon_{it}$$

using data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for borrowers born between 1930 and 1953. In Panel A, *riskscore* is a proprietary metric from Equifax that measures a consumer's creditworthiness. It ranges from 280 to 850. In Panel B, the dependent variable *frac accounts delinquent* equals the number of accounts at least 90 days past due divided by the total number of credit accounts. *Tribalcourt* equals one if the consumer's first credit report is on a reservation using tribal courts as determined by Public Law 280. *Median income* comes from the 2000 US census and is at the census tract level. Standard errors are clustered by census tract and date. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

Panel A: Financial health measured by <i>riskscore</i>			
	(1a)	(2a)	(3a)
<i>tribalcourt</i>	−1.471 (3.93)	3.359 (3.81)	3.317 (3.81)
<i>median income / 1000</i>	0.941*** (0.13)	0.929*** (0.14)	0.931*** (0.14)
birth year FE	x	x	x
age FE	x	x	x
date quarter FE	x	x	
census region FE		x	
census region × date quarter FE			x
<i>N</i>	175,970	175,970	175,970
<i>R</i> ²	0.100	0.120	0.120
Panel B: Financial health measured by <i>frac. accounts delinquent</i>			
	(1b)	(2b)	(3b)
<i>tribalcourt</i>	0.00355 (0.0048)	−0.00287 (0.0041)	−0.00284 (0.0041)
<i>median income / 1000</i>	−0.000696*** (0.00015)	−0.000703*** (0.00017)	−0.000703*** (0.00017)
birth year FE	x	x	x
age FE	x	x	x
date quarter FE	x	x	
census region FE		x	
census region × date quarter FE			x
<i>N</i>	148,109	148,109	148,109
<i>R</i> ²	0.0085	0.016	0.020

(employment or poverty rate), education levels, marriage rate, or the credit profile of older cohorts, *tribalcourt* has a significant negative effect on consumer riskscores and, in all but one case, a significant positive effect on delinquencies.¹⁴

We complement our main tests by performing a placebo exercise using individuals who were born prior to the passage of PL280 but who currently reside on a reservation—birth cohorts from 1930 to 1953. Because this older cohort grew up before PL280's effects, this sample is useful to distinguish the effects of formative exposures from current financial markets. The results in Table 5 show that the financial health of these older generations is similar on tribal and state court reservations, which suggests that different formative experiences (absent for this older cohort) are important for consumer financial health.

Fig. 6 presents how differences between tribal and state court consumers depend on the length of formative exposure to differences in local banking by examining the effect for different age cohorts. Younger cohorts were exposed to differences in local banking throughout their formative years, whereas older cohorts grew up when there were no striking differences between tribal and state court reservations. Consistent with the timing of the financial development differences between tribal and state court reservations, we find that the difference in credit scores is greatest among younger cohorts and declines in the amount of formative exposure. Tribal court consumers have worse credit scores than state court consumers for all cohorts younger than the 50-year-old age cohort. Indeed, given the sharp differences in current financial conditions facing these consumers, the results in Table 5 and Fig. 6 suggest that formative conditions are critical for subsequent financial health. In addition, these findings rule out a number of alternative explanations for the poor financial health of young adult tribal court residents. For example, most omitted variables affecting the quality and demand for credit in local reservation areas would also affect the financial health of the older generation, and thus the null findings on the financial health of the older generation provide compelling evidence against a wide variety of alternative interpretations.

¹⁴ The one variable that, when included as a control variable, makes the coefficient on *tribalcourt* lose statistical significance is the average number of credit lines for reservation area borrowers who are born before PL280's passage. This variable potentially proxies for the level of financial development on reservations, which could explain why *tribalcourt*—our preferred measure of financial development across reservations—loses significance when both variables are on the regression's right hand side.

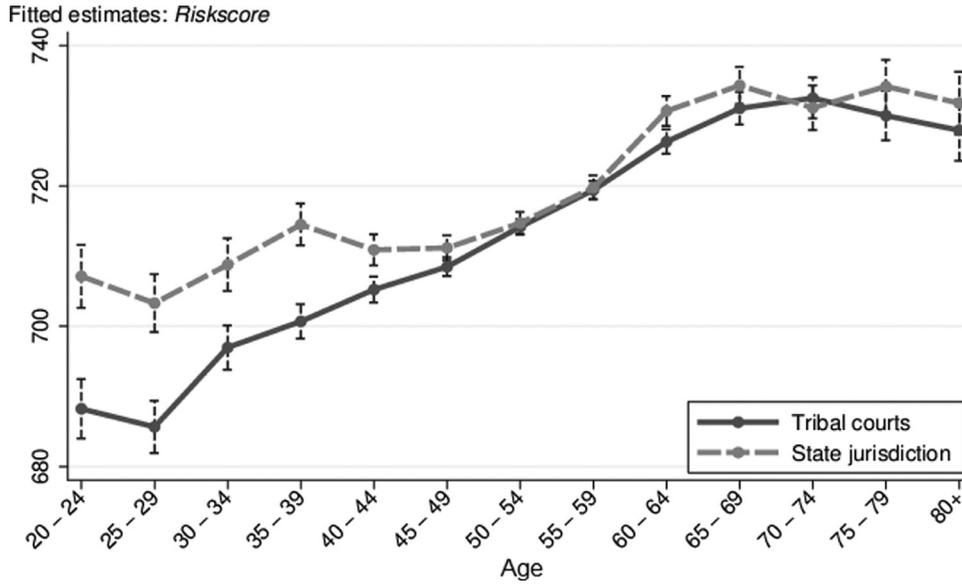


Fig. 6. Credit scores by age cohorts.

This figure uses data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for borrowers who would have turned 18 by 1999 and whose first credit report is associated with an address on Native American reservation lands as defined by the Bureau of Indian Affairs. The sample contains consumers who are only on reservation lands during the sample. The figure illustrates the fitted model,

$$riskscore_{it} = \gamma_t + \gamma_{birthyear} + \gamma_{age} + \beta_1 tribalcourt_{it} + \sum_k \beta_{2k} \cdot D_k^{age} + \sum_k \beta_{3k} \cdot D_k^{age \cdot tribalcourt} + \varepsilon_{it}.$$

The figure includes 95% prediction intervals calculated using standard errors clustered by census tract and date.

5.2. Evidence from movement away from reservation areas

Next, we study the dynamics of financial health for consumers who move away from reservation areas. This test quantifies how well exposure to stronger financial environments mitigates the early life exposure effects we observed in the main tests. Specifically, we consider a difference-in-differences specification that compares improvements to financial health for consumers departing tribal court reservations and improvements to financial health for consumers departing state court reservations. Because the identifying variation in our tests is over time, these specifications allow us to include more granular fixed effects for the consumer’s first census tract to address any lingering concerns over unobserved differences in borrower quality or economic opportunity.

5.2.1. Specification for moving away from reservations

We test the effects of moving away from reservation areas by estimating the following difference-in-difference regression:

$$Y_{it} = \gamma_t + \gamma_{birthyear} + \gamma_{age} + \gamma_{age \text{ when move}} + \gamma_{state \text{ move to}} + \gamma_{first \text{ census tract}} + \beta_1 years \text{ away}_{it} + \beta_2 years \text{ away}_{it} \cdot tribalcourt_i + \beta_3 income_{it} + \beta_4 riskscore \text{ when move}_i + \beta_5 years \text{ away}_{it} \cdot riskscore \text{ when move}_i + \varepsilon_{it}, \quad (5)$$

where Y_{it} measures consumer credit outcomes, and $yearsaway$ equals the number of years since consumer i moved off of reservation lands (equals zero for on-reservation observations). Fixed effects are for quarter t

(γ_t), the year of consumer i ’s birth ($\gamma_{birthyear}$), consumer i ’s age in quarter t (γ_{age}), the age when consumer i moves off the reservation ($\gamma_{age \text{ when move}}$), and the census tract of i ’s first credit record ($\gamma_{first \text{ census tract}}$). Some specifications also include fixed effects for i ’s current state of residence (γ_s). All specifications control for the median level of *income* on the census tract where consumer i is located at time t as well as the consumer’s *riskscore* at the time they move off of the reservation. Controlling for *riskscore when move* and its interaction with *yearsaway_{it}* is a flexible way to account for potential differences in the quality of borrowers who leave state versus tribal court reservations.¹⁵

The coefficient β_1 captures the baseline effect of moving away from reservations with state courts, whereas β_2 measures the differential change in credit outcomes for individuals moving away from tribal court reservations. By capturing this differential effect of *yearsaway* for consumers from tribal court reservations, the estimate of β_2 indicates how quickly credit scores of individuals from tribal court reservations converge with individuals

¹⁵ The evidence in online Internet Appendix Table A.5 shows that borrowers from tribal court reservations are less likely to move compared to individuals from state court reservations. To the extent these findings reflect constraints that disproportionately prevent individuals from moving from tribal court reservations, our main estimates understate how much the financial health of the typical resident of a tribal court reservation would change if they left the reservation. It is also possible that the constraints on leaving tribal court areas lead to systematic differences the types of individuals who select out-of-state versus tribal reservations. Our specification addresses this concern by directly controlling for each individual’s financial health at the time they move and interacting those initial conditions with *years away*.

from state court reservations. Expressing the difference-in-difference in yearly terms also helps to account for the influence of post-move truncation of the sample, placing the post-move observations on equal footing whether they come from an individual who has two years of data post move or one who has ten years of data post move.¹⁶ In addition, the term *years away · riskscore when move* is included to ensure that the estimated coefficient β_2 is not merely reflecting a natural convergence to the mean in which individuals with low credit scores naturally improve more over time. The inclusion of this term is not essential to our findings but is useful to show robustness.

An advantage of the specification in Eq. (5) is the richness of the fixed effects, which account flexibly for unobserved geographical variation in economic activity. For example, the model compares two consumers, one from a tribal court area and one from a state court reservation, both of whom move to the same state. Because the model has fixed effects for *i*'s current state, it accounts for arbitrary differences in the broader economic activity of the destination area where the consumer moves. Further, because there are at least two consumers who originate from the same census tract (some of whom stay on reservations and others that leave), the specification establishes a baseline effect for the economic conditions when *i* first establishes his or her credit report. The primary source of variation that remains is the plausibly exogenous difference in local financial markets across state and tribal court jurisdictions.

5.2.2. Regression estimates of moving away from reservations

The financial health benefits of moving away from a reservation are stronger for consumers who come from tribal court reservations. Table 6, Panel A presents estimates of Eq. (5) using *riskscore* as the dependent variable. The coefficient of interest is β_2 , which indicates whether years away from the reservation has a relatively stronger effect on credit scores of individuals who move from tribal court reservations compared to individuals who move from state court reservations. Consistent with financial underdevelopment stunting credit records of borrowers growing up on tribal court reservations, the estimated coefficient on the interaction term is positive and statistically significant at the 1% level in all specifications. Our strongest evidence comes from the specification that includes fixed effects for the consumer's state of residence at the current quarter *t* (column 4), which broadly accounts for the economic conditions and quality of financial markets of the postreservation destination.¹⁷ The coefficient estimates are robust to the inclusion of these fixed effects, and the estimates indicate large differential effects for tribal court movers. In particular, for consumers leaving state court areas, risk

scores increase by just over 1 point for each year they are away from the reservation, whereas the corresponding yearly increase is approximately 1.6 points (38% greater) for consumers leaving tribal reservations.

Consumers who move away from tribal court reservations also have a relatively larger reduction in the fraction of delinquent accounts. We estimate Eq. (5) using *frac. accounts delinquent* as the dependent variable (Table 6, Panel B). Using the same set of fixed effects as Panel A, the estimate of β_2 is approximately -0.001 and is statistically significant at the 1% level across specifications. Notably, in the specifications with fixed effects for age when moved and current state of residence, the baseline effect of state court movers is negative but very small in magnitude and not statistically different from zero. These estimates imply that moving away from state court reservations has little impact on the way consumers manage credit, while leaving tribal court reservations leads to a gradual decline in the fraction of delinquent accounts, consistent with a causal connection between the local financial market consumers are exposed to and their financial health.

5.2.3. The long-run erosion of exposure effects

As an additional consideration, we use the estimates in Table 6 to examine whether movement to areas with stronger local financial markets can overcome the negative effect that growing up in areas with low financial development has on consumer financial health. The exercise is useful because it empirically distinguishes the effects of formative experiences from the effects of current credit market conditions.

The estimates in Table 6 suggest that exposure to stronger local financial markets partially offsets the effect of early exposure to areas with less financial development, but the consequences of early life experiences take a long time to overcome. Fig. 7 presents fitted estimates of Eq. (5). In the top panel, the increase in risk scores is faster for consumers from tribal court reservations, but these consumers have worse financial health when they leave the reservation (*riskscore* equal to 634 versus 646 when *years away* = 0 for tribal and state court, respectively). It takes almost 17 years for the average *riskscore*s of tribal court and state court reservation movers to no longer be statistically different from each other. Similarly, the fraction of delinquent accounts falls much faster for consumers leaving tribal court reservations, but it takes approximately 12 years before average delinquency rates are no longer statistically different (and this relatively faster rate of convergence in the bottom panel is due to wider prediction intervals around the estimated effects). These results are a strong indication that early exposure to financial markets is an important determinant of consumer financial health that is not easily transformed by later experiences. In this way, the estimates suggest that interventions that affect formative experiences have long-lived effects on household financial health.

6. Robustness and mechanisms

In this section, we describe a series of tests that provide insight into the mechanisms through which early life

¹⁶ We obtain qualitatively similar estimates in specifications where we employ a post-move indicator, but we prefer the *years away* specification because of the potential issue with truncation and the information it provides about persistent differences in financial health after leaving tribal court reservations.

¹⁷ Because the consumers in our sample geographically spread out upon moving from the reservation, state fixed effects are the most granular destination fixed effect we can employ without losing the ability to identify the interaction between *years away* and *tribalcourt*.

Table 6

Moving away from reservations and consumer creditworthiness.

This table presents OLS estimation results of the following specification,

$$Y_{it} = \gamma_t + \gamma_{\text{birthyear}} + \gamma_{\text{age}} + \gamma_{\text{age when move}} + \gamma_{\text{state move to}} + \gamma_{\text{first census tract}} + \beta_1 \text{years away}_{it} + \beta_2 \text{years away}_{it} \cdot \text{tribalcourt}_i + \beta_3 \text{income}_{it} + \beta_4 \text{riskscore when move}_i + \beta_5 \text{years away}_{it} \cdot \text{riskscore when move}_i + \varepsilon_{it},$$

using data from FRBNY-CCP, a 5% random panel of consumer credit reports from Equifax. The sample includes quarterly observations on a panel of consumer credit records between 1999Q1 and 2015Q2 for borrowers who have turned 18 by 1999 and whose first credit report is associated with an address on Native American reservation lands as defined by the Bureau of Indian Affairs. In Panel A, *riskscore* is a proprietary metric from Equifax that measures a consumer’s credit-worthiness. It ranges from 280 to 850. In Panel B, *frac accounts delinquent* equals the number of accounts at least 90 days past due divided by the total number of credit accounts. *Years away from resvn* is the number of years that have passed since *i* has moved off of the reservation lands. $\gamma_{\text{first census tract}}$ is a fixed effect for the Census tract of *i*’s first credit report. $\gamma_{\text{age when move}}$ is a fixed effect for *i*’s age when they leave the reservation that is set to zero for all *i* that do not leave the reservation during the sample. $\gamma_{\text{state move to}}$ is a fixed effect for the state that *i* moves to when they leave the reservation that is set to zero for all *i* that do not leave the reservation during the sample. *Tribalcourt* equals one if the consumer’s first credit report is on a reservation using tribal courts as determined by Public Law 280. *Median income* comes from the 2000 US Census and is at the census tract level. The regressions also control for *i*’s *riskscore* when they leave the reservation and its interaction with *years away from resvn*. Standard errors are clustered by census tract and date. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

Panel A: Financial health measured by riskscore			
	(1a)	(2a)	(3a)
<i>tribalcourt</i> × <i>years away from resvn</i>	0.401*** (0.086)	0.411*** (0.087)	0.457*** (0.087)
<i>years away from resvn</i>	1.987*** (0.083)	1.192*** (0.093)	1.178*** (0.093)
<i>median income / 1000</i>	0.247*** (0.011)	0.248*** (0.011)	0.250*** (0.011)
birth year FE	x	x	x
age FE	x	x	x
date quarter FE	x	x	x
first census tract FE	x	x	x
age when move FE		x	x
state moved to FE			x
N	349,445	349,445	349,445
R ²	0.51	0.51	0.52
Panel B: Financial health measured by frac. accounts delinquent			
	(1b)	(2b)	(3b)
<i>tribalcourt</i>	−0.00960*** (0.00034)	−0.000903*** (0.00033)	−0.00113*** (0.00034)
<i>years away from resvn</i>	−0.00742*** (0.00036)	−0.000193 (0.00042)	−0.0000967 (0.00042)
<i>median income / 1000</i>	−0.000605*** (0.000042)	−0.000625*** (0.000044)	−0.000600*** (0.000042)
birth year FE	x	x	x
age FE	x	x	x
date quarter FE	x	x	x
first census tract FE	x	x	x
age when move FE		x	x
state moved to FE			x
N	284,722	284,722	284,722
R ²	0.21	0.21	0.22

exposure to banking improves consumer credit profiles. Our tests complement the main analysis of the FRBNY-CCP data with evidence from the Qualtrics survey and other new data sets matched to the reservation setting. In these tests, we find consistent evidence that early life exposure to finance improves financial literacy (particularly relating to working knowledge of retail banking products), and enhances trust in financial institutions.

6.1. Financial literacy and trust

A possible mechanism for our findings is that early life exposure to finance confers experience that substantively improves financial literacy, which leads to long-term improvements in consumer credit profiles. We present a series of tests to evaluate this financial literacy channel.

Our first test exploits state-by-state variation in financial education requirements to graduate from secondary school. Specifically, Brown et al. (2016) analyze a series of state-level schooling reforms that increased graduation requirements in financial, economics, and mathematics during our sample period. Using the FRBNY-CCP data, we examine whether a financial education mandate offsets the effect of growing up without finance. This test augments the specification in Eq. (2)—how long it takes after turning 18 years old for consumers to get their first credit report or line of credit—with the interaction between *tribalcourt* and an indicator for needing financial education training to graduate. According to Table 7, the interaction effect is positive and significant across specifications. Indeed, the financial education mandates have no impact on credit market entry for individuals growing up on state court

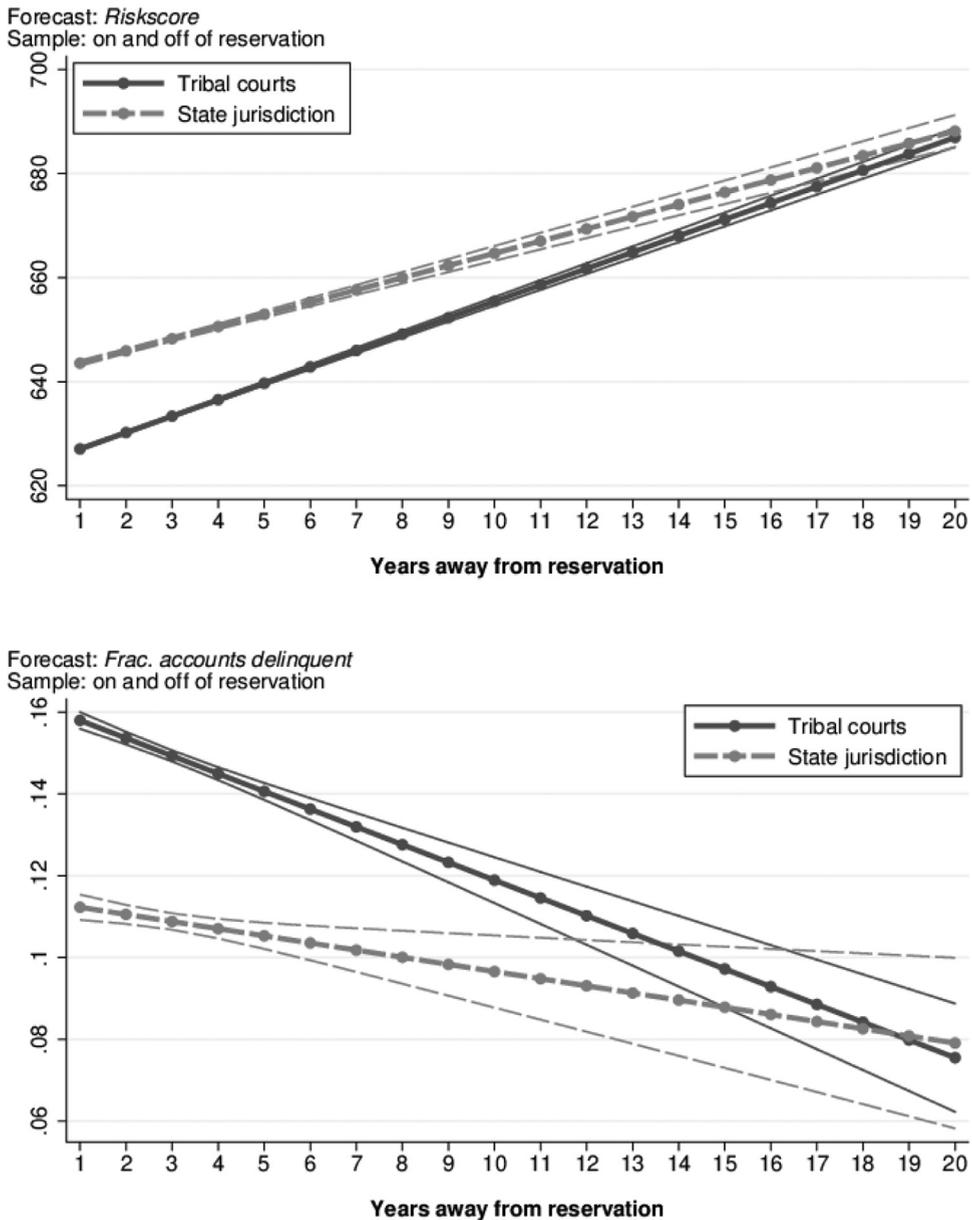


Fig. 7. Financial health after moving from reservations. This figure presents fitted estimates of the regression,

$$Y_{it} = \gamma_1 + \gamma_{birthyear} + \gamma_{age} + \gamma_{age\ when\ move} + \gamma_{state\ move\ to} + \gamma_{first\ census\ tract} + \beta_1 years\ away_{it} + \beta_2 years\ away_{it} \cdot tribalcourt_i + \beta_3 income_{it} + \varepsilon_{it},$$

described in Table 6. The figure represents 95% prediction intervals calculated using standard errors clustered by date and the census tract of *i*'s first credit report.

reservations but completely offset the negative effects of growing up without finance for tribal court residents. By contrast, the specifications in columns 2, 3, 5 and 6 show that analogous general economics and math education requirements do not offset the credit market disadvantages of growing up without finance. These findings suggest that the adverse consequences of growing up without finance can be alleviated by improvements to financial education but not by general improvements to education. Moreover, because these education reforms are unrelated to changes in economic opportunity, the pattern of results suggests

the effects of growing up without finance operate through a financial literacy channel that is distinct from other mechanisms.

Our second test utilizes data from the CFPB's consumer complaints database to provide additional evidence on how growing up without finance affects consumers' understanding of credit market information. We use the ZIP code of each consumer complaint in the CFPB database to match the data to the reservation setting, analogous to the FRBNY-CCP's geographic match with reservation lands. Using the matched data, we empirically assess whether a

Table 7

Credit market entry and financial education.

This table presents linear probability estimates using the same samples and specifications as Table 2. This table uses the state-year-level public education reforms from Brown, Grisby, van der Klaauw, Wen, and Zafar (2016). The education variables—financial literacy mandate, economics mandate, and mathematics reform—equal one for consumers in a given state who would have graduated after the new policies were implemented. Standard errors are clustered by census tract and date. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

	time to first credit report			time to first line of credit		
	(1)	(2)	(3)	(4)	(5)	(6)
<i>tribalcourt</i> × <i>financial literacy mandate</i>	0.0429* (0.024)	0.0401* (0.024)	0.0460* (0.025)	0.0243** (0.011)	0.0252** (0.011)	0.0239** (0.010)
<i>financial literacy mandate</i>	-0.00819 (0.023)	-0.00631 (0.023)	-0.0126 (0.024)	-0.0127 (0.010)	-0.0133 (0.010)	-0.0121 (0.010)
<i>tribalcourt</i> × <i>economics mandate</i>		0.00772 (0.013)			0.00303 (0.0080)	
<i>economics mandate</i>		0.00834 (0.014)			-0.00851 (0.0081)	
<i>tribalcourt</i> × <i>mathematics reform</i>			-0.00682 (0.015)			0.00137 (0.0091)
<i>mathematics reform</i>			0.0120 (0.015)			-0.00195 (0.0086)
median incomes	x	x	x	x	x	x
date quarter FE	x	x	x	x	x	x
age FE	x	x	x	x	x	x
state FE	x	x	x	x	x	x
<i>N</i> (consumer quarter)	151,394	151,394	151,394	246,735	246,735	246,735
<i>N</i> (consumers)	14,380	14,380	14,380	14,380	14,380	14,380
<i>R</i> ²	0.12	0.12	0.12	0.048	0.048	0.048

Table 8

Consumer complaints and recourse.

This table uses data from the Consumer Financial Protection Bureau’s database of consumer complaints matched to ZIP codes on Native American reservations. It presents OLS estimation results of the following specification,

$$credit\ outcome_i = \beta_1 tribalcourt_i + \beta_2 controls_i + \varepsilon_i.$$

In columns 1–3, the dependent variable equals one if the complaint claims incorrect information on a credit report. In columns 4–6, the dependent variable equals one if the complaint was resolved without monetary relief. Standard errors are clustered by state. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

sample: dependent variable:	all complaints credit information issue (= 1)			has credit information issue no relief on complaint (= 1)		
	(1)	(2)	(3)	(4)	(5)	(6)
<i>tribalcourt</i>	0.0544*** (0.015)	0.0694*** (0.018)	0.0684*** (0.021)	0.0962* (0.056)	0.0816* (0.043)	0.0819* (0.047)
personal incomes (ZIP code)			0.00342* (0.0018)			-0.00956 (0.0056)
Census region FE		x	x		x	x
Number of complaints	1,604	1,604	1,568	197	197	197
<i>R</i> ²	0.0061	0.013	0.017	0.0083	0.033	0.051

consumer’s complaint is a credit information issue (identified by the “Issue” field in the CFPB database) as well as whether the complaint was resolved without relief (either “resolved with explanation” or “resolved with no relief”). We find that individuals from tribal court areas are more likely to report credit information issues (Table 8, columns 1 through 3), but conditional on reporting a credit information issue, they are substantially less likely to receive relief for their complaint (columns 4 through 6). This evidence suggests individuals who grow up without finance have less understanding of credit market information, which is consistent with lower financial literacy.

Our third test uses responses to our custom Qualtrics survey of Native Americans who grew up near reservation lands. Specifically, we asked respondents to answer the “Big Five” financial literacy questions, whose language we

take directly from Lusardi and Mitchell (2014). The questions test the respondents understanding of compound interest, inflation, stock portfolio diversification, bond pricing, and mortgage payments. Clearly, some of these questions are more likely to be affected by exposure to retail banking than others. In particular, questions about stock diversification and bond pricing are unlikely to be aspects of literacy that an individual would learn from experience with a retail bank. On this intuition, we construct two separate financial literacy indexes: one for bank-centric literacy (i.e., compound interest, inflation, and mortgage payments) and another for stock/bond literacy (i.e., stock portfolio diversification and bond pricing). We relate these separate financial literacy measures to whether an individual reports having grown up in a community with a bank.

Table 9

Financial literacy and trust, survey evidence.

This table uses data from a survey administered by Qualtrics. The survey was targeted to respondents 40 years old or younger who identify as Native American and grew up on or near a Native American reservation. This table presents OLS estimation results of the following specification,

$$\text{personal attribute}_i = \beta_1 \text{grow up near bank}_i + \beta_2 \text{controls}_i + \varepsilon_i.$$

The dependent variable in columns 1 and 2 is equal to the number of correct answers to the following three banking-centric financial literacy questions: "Suppose you had \$100 in a savings account and the interest rate was 2% per year. After 5 years, how much do you think you would have if you left the money to grow?" More than \$102, Exactly \$102, Less than \$102, Do not know, Refuse to answer. "Imagine that the interest rate on your savings account by 1% per year and inflation was 2% per year. After 1 year, how much would you be able to buy with the money in this account?" More than today, Exactly the same, Less than today, Do not know, Refuse to answer. "A 15-year mortgage typically requires higher monthly payments than a 30 year mortgage, but the total interest over the life of the loan will be less." True, False, Do not know, Refuse to answer. In columns 3 and 4, the dependent variable equals the number of correct answers to the following stock/bond literacy questions: "A 15-year mortgage typically requires higher monthly payments than a 30 year mortgage, but the total interest over the life of the loan will be less." True, False, Do not know, Refuse to answer. "Buying a single company's stock usually provides safer return than a stock mutual fund." True, False, Do not know, Refuse to answer. In columns 5 and 6, respondents rate on a scale of 1 to 7, "How much do you trust bank officials to help with your financial decisions?" In columns 7 and 8, the dependent variable equals one if the respondent answers "Most people can be trusted" to the question "Generally speaking, would you say that most people can be trusted or that you have to be very careful in dealing with people?" The independent variable of interest, *grow up near bank*, equals one if the respondents answers yes to the question "Was there a bank in the community where you grew up?" Standard errors are clustered by state. Stars *, **, and *** indicate statistical significance at the 10%, 5%, and 1% levels.

dependent variable:	bank-centric fin. lit.		stock/bond fin. lit.		financial trust		general trust	
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
<i>grow up near a bank</i>	0.231*** (0.065)	0.171*** (0.061)	-0.0917 (0.073)	-0.0952 (0.076)	0.596*** (0.18)	0.558*** (0.20)	-0.0127 (0.024)	-0.0245 (0.027)
<i>male</i>	0.179* (0.098)	0.165 (0.11)	0.0432 (0.059)	0.0440 (0.057)	0.158 (0.16)	0.115 (0.16)	0.0200 (0.029)	0.0215 (0.031)
<i>age</i>	-0.0859 (0.059)	-0.0858 (0.063)	0.0296 (0.043)	0.0223 (0.044)	-0.0258 (0.096)	0.00607 (0.097)	0.0361** (0.018)	0.0322* (0.018)
<i>age squared</i>	0.00173 (0.0010)	0.00170 (0.0011)	-0.000458 (0.00079)	-0.000346 (0.00080)	0.000422 (0.0017)	-0.000111 (0.0017)	0.000613* (0.00031)	0.000543* (0.00031)
Census region FE	x	x	x	x	x	x	x	x
Income FE		x		x		x		x
Education FE		x		x		x		x
Number of complaints	704	704	704	704	704	704	661	661
R ²	0.060	0.098	0.026	0.054	0.039	0.068	0.025	0.045

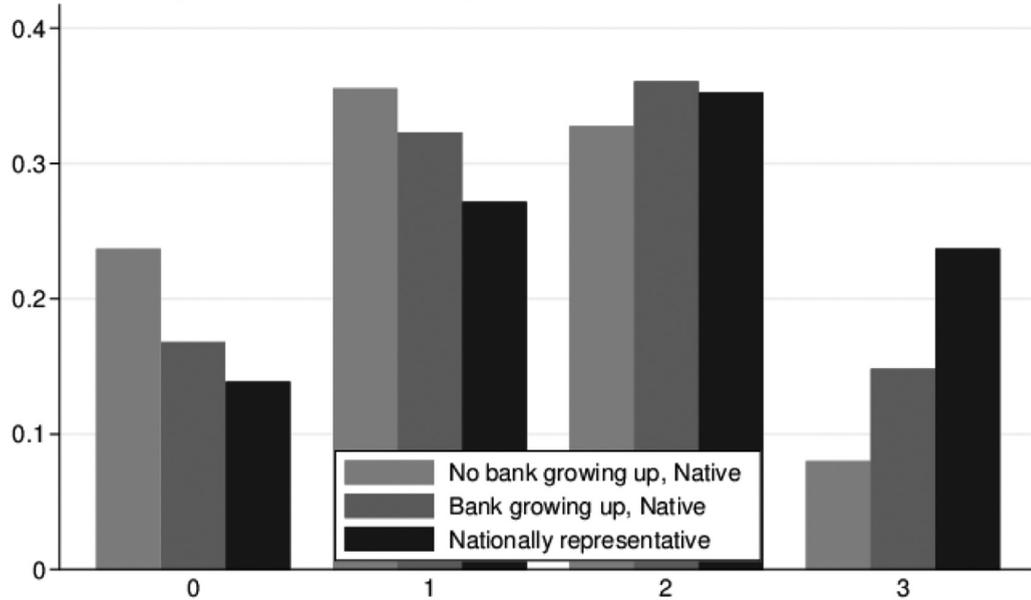
The first panel of Fig. 8 presents the distribution of bank-centric financial literacy, separately for individuals who grew up without finance, individuals who grew up with finance, and, as a comparison, our nationally representative subsample. Consistent with a financial literacy channel, individuals without formative exposure to banking have lower bank-centric financial literacy than individuals who grew up near a bank. Further, in columns 1 through 4 of Table 9, we evaluate this relation in a regression that controls for gender, age (and age squared), census region fixed effects, and individual-level bins for income and education. Consistent with a financial literacy mechanism whereby exposure to banks improves financial knowledge related to retail banking, we find a strong positive relation between bank-specific financial literacy and growing up near a bank, but we find no indication that growing up near a bank improves stock/bond financial literacy. Apart from corroborating our evidence on the effects of financial (versus economics/math) education reforms, the survey evidence provides precise insight into the specific channel through which early life exposure to banking improves consumer credit profiles. In particular, the survey suggests that these effects are not just a general improvement to knowledge but specific improvements related to the nature of banking services.

A related mechanism that is consistent with the financial literacy channel is that growing up without finance affects trust in financial institutions. Beyond substantive

skills, exposure to finance might make individuals more trusting of financial institutions and thus willing to engage with credit markets at an earlier age. To evaluate whether growing up without finance affects trust in financial institutions, we turn to the survey in which we ask individuals about the degree to which they trust advice from bankers, as well as their generic levels of trust in other individuals. As with the financial literacy indexes, we relate these measures of trust to whether an individual grew up in a community with a bank.

The second panel of Fig. 8 presents the distribution of finance-specific trust from our survey, separately for individuals who did not grow up near a bank, individuals who grew up near a bank, and our nationally representative subsample. Consistent with a financial trust mechanism, the distribution of financial trust is shifted to the right for individuals who grew up in a community with a bank. We further subject the relation between growing up near a bank and financial trust to a regression analysis with controls and fixed effects. As columns 5 through 8 of Table 9 show, there is a strong and significant positive relation between growing up near a bank and the financial trust score (out of seven) that is robust to individual-level controls for education and income as well as census region fixed effects. On the other hand, there is no relation between growing up without finance and generalized trust (columns 5 and 6 of Table 9). Analogous to our findings on bank-specific financial literacy, this finding is a strong

Financial literacy test score (fraction respondents)



How much trust bank officials' help w/ financial decisions? (fraction respondents)

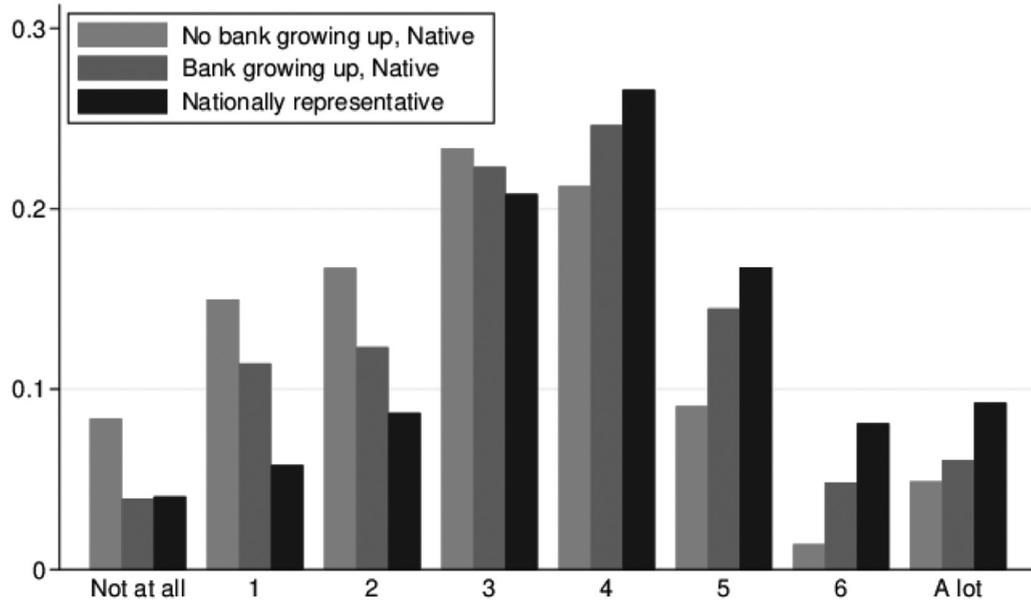


Fig. 8. Surveyed financial trust and financial literacy.

This figure uses data from a novel survey administered by Qualtrics. The survey was targeted to Native American individuals younger than forty years old who grew up on or near reservation lands. The figure also includes a sample that is nationally representative in terms of race. Financial literacy test score is on a scale of 0 to 3, with 3 meaning that the respondent answers all of the following multiple choice questions correctly: "Suppose you had \$100 in a savings account and the interest rate was 2% per year. After 5 years, how much do you think you would have if you left the money to grow?" More than \$102, Exactly \$102, Less than \$102, Do not know, Refuse to answer. "Imagine that the interest rate on your savings account by 1% per year and inflation was 2% per year. After 1 year, how much would you be able to buy with the money in this account?" More than today, Exactly the same, Less than today, Do not know, Refuse to answer. "A 15-year mortgage typically requires higher monthly payments than a 30 year mortgage, but the total interest over the life of the loan will be less." True, False, Do not know, Refuse to answer. To measure financial trust, the survey asks respondents to rate on a scale of 1 to 7, "How much do you trust bank officials to help with your financial decisions?"

indication that experience with banks is indeed responsible for the differences in trust.

6.2. Economic opportunity and other mechanisms

Economic opportunity is certainly an important determinant of individuals' long-run credit outcomes. However, even after controlling for incomes in Tables 3 and 4, our results show that early life exposure to finance is also an important determinant of consumer credit profiles and is distinct from different economic opportunities. Identifying a role for nonincome factors that improve consumer credit profiles is important for policy-making, particularly given the difficulty implementing policies which broadly improve economic opportunities.¹⁸

Several aspects of our tests indicate that the effects of growing up without finance are not entirely due to differences in economic opportunity: (1) in each of our tests, we control for income at the most granular level possible, finding consistent results when we can control for income at the individual level in the survey (Table 3); (2) beyond income, demographic controls that reflect economic opportunity have little impact on our estimates (Fig. 5); (3) shocks to bank branching and financial education, which are unrelated to economic development or PL280 status, mitigate the effects of local financial development, suggesting these factors have an independent effect on consumer financial health; (4) the effect of growing up without finance is absent for individuals who were born before the differences in financial development emerged and are smaller than our main estimates for middle-aged individuals who had less exposure to these differences (Table 5 and Fig. 6); and (5) the effects of early life experiences are persistent, even for individuals who move from different formative environments to the same current environment (Table 6 and Fig. 7).

Furthermore, we explore other mechanisms that, according to standard economic theories of consumer decision-making, would suppress credit market entry and lead to worse credit outcomes after obtaining credit. First, we consider the role of attitudes toward risk. Growing up in a community with a bank is not associated with any difference in survey-measured risk aversion or ambiguity aversion (see online Internet Appendix Table A.3 for these and other results from the survey). In addition, using proprietary data on cash access at casinos, we find that state and tribal court residents have similar levels of spending at casinos (online Internet Appendix A.5).¹⁹ Taken together, these results suggest that individuals from state and tribal

court reservations have similar attitudes toward risk. Second, individuals who grow up without finance could have less confidence in their financial aptitude, but the survey responses indicate that consumers from communities with and without banks have similar levels of financial confidence. Finally, our findings could relate to differences in discount rates. Using the survey, we find that individuals who grow up near a bank consider themselves to be more patient, which would translate into less demand for consumption today relative to tomorrow. This can potentially explain why tribal court residents have higher levels of delinquencies upon obtaining credit. However, we would expect that tribal court residents, if they are less patient, would have greater demand for credit and therefore show up in the credit bureau data at a younger age. Hence, differences in discount rates are difficult to reconcile with our findings.²⁰

Overall, our findings speak to a particular mechanism through which early life exposure to banking improves long-run consumer credit profiles. In light of this evidence, our results suggest that encouraging more bank branches and improving financial literacy and trust would improve consumer financial health. Indeed, the effects of these policy responses can be seen in Tables 2 and 7, which show that banking deregulation (encouraging banking development) and financial literacy reforms each offset the effect of growing up without finance on early access to credit. Of course, local economic development would likely also improve consumer credit profiles, but the independent effect of financial literacy and financial trust on financial health is an important finding, particularly in a policy environment in which job creation is difficult. Given the challenges in creating jobs on reservations and the importance of household financial health, these additional policy measures are potentially quite valuable.

7. Conclusion

This paper shows that formative exposure to local financial markets have large, persistent effects on how an individual builds and manages credit over a lifetime. Our approach marries location-specific micro-level data on consumer financial health with large and unintended differences in financial development across Native American reservations arising from US Congressional action in 1953. We find that individuals growing up in areas with relatively strong financial markets establish a credit history sooner and have better financial health in terms of both higher credit scores and less frequent credit account delinquencies. Moreover, although individuals who leave areas with weak financial markets see significant improvements

¹⁸ Improving economic opportunities in areas with persistent underdevelopment, such as Native American reservations, is especially challenging. For example, see testimony in the Congressional hearing "Unemployment on Indian reservations at 50 percent: the urgent need to create jobs in Indian country" Hearing before the Committee on Indian Affairs, US Senate, One Hundred Eleventh Congress, Second Session, January 28, 2010.

¹⁹ The casino data set contains information on the residential location of casino patrons (ZIP code) for a broad cross-section of U.S. casinos. Using the information on residential location of casino patrons, we restrict our tests of gambling consumption to individuals who live on a reservation, but the gambling expenditure can occur at casinos across the United States. Accordingly, this test does not reflect whether there are more off-

reservation visitors to a tribal casino. For more detail on the data set, see Cookson (2018) who uses a distinct subsample of the national data set to study the introduction of prize-linked savings in Nebraska.

²⁰ We also provide evidence that suggest our results are not merely due to differences in credit demand. After controlling for income, consumers on state and tribal court reservations have similar numbers of hard credit inquiries (online Internet Appendix Table A.2). Should tribal court status merely proxy for unobserved differences in economic opportunity, we would expect tribal court residents to have less credit demand.

in consumer financial health, it takes many years to overcome the negative effects of growing up without finance.

These findings provide new insights into the consumer-side effects of financial development and, in particular, highlight important long-run consequences of local financial market development for household well-being. In this way, our work not only speaks to the long-term benefits of financial inclusion, but also suggests that traditional banking institutions matter through an underappreciated channel—early life exposure to financial markets, which enhances financial literacy and trust in financial institutions. This insight is important to consider as traditional local financial institutions continue to consolidate and move services online.²¹ Although financial institutions appear to be ubiquitous, there remain important gaps in local financial development beyond Native American reservations—for example, see McDevitt and Sojourner (2016)'s example of fringe banking in the Bronx, New York. By showing that these gaps have economically large effects on long-term household financial health, our findings suggest that much more work is needed to understand how these gaps form in the first place and to study effective policies to remedy them.

References

Addoum, J.M., Korniotis, G., Kumar, A., 2016. Stature, obesity, and portfolio choice. *Manag. Sci.* 63 (10), 3393–3413.

Agarwal, S., Chomsisengphet, S., Mahoney, N., Stroebel, J., 2015. Regulating consumer financial products: Evidence from credit cards. *Q. J. Econ.* 130 (1), 111–164.

Agarwal, S., Driscoll, J.C., Gabaix, X., Laibson, D., 2009. The age of reason: Financial decisions over the life cycle and implications. *Brook. Pap. Econ. Act.* Fall, 51–117.

Anagol, S., Balasubramaniam, V., Ramadorai, T., 2015. The effects of experience on investor behavior: evidence from India's IPO lotteries. Working Paper, SSRN, 2568758

Anderson, T.L., Parker, D.P., 2008. Sovereignty, credible commitments, and the prosperity of American Indian reservations. *J. Law Econ.* 51 (4), 641–666.

Anderson, T.L., Parker, D.P., 2017. *The Law and Economics of Federalism. Lessons in Fiscal Federalism from American Indian Nations.* Edward Elgar Publishing, pp. 55–90.

Baker, S.R., 2018. Debt and the response to household income shocks: validation and application of linked financial account data. *J. Polit. Econ.* 126 (4), 1504–1557.

Barnea, A., Cronqvist, H., Siegel, S., 2010. Nature or nurture: what determines investor behavior? *J. Financ. Econ.* 98 (3), 583–604.

Beck, T., Demirguc-Kunt, A., Peria, M., 2007. Reaching out: access to and use of banking services across countries. *J. Financ. Econ.* 85, 234–266.

Berkowitz, D., Lin, C., Ma, Y., 2015. Do property rights matter? Evidence from a property law enactment. *J. Financ. Econ.* 116 (3), 583–593.

Brown, J.R., Cookson, J.A., Heimer, R., 2017. Law and finance matter: lessons from externally imposed courts. *Rev. Financ. Stud.* 30 (3), 1019–1051.

Brown, M., Grigsby, J., van der Klaauw, W., Wen, J., Zafar, B., 2016. Financial education and the debt behavior of the young. *Rev. Financ. Stud.* 29 (9), 2490–2522.

Butler, A.W., Cornaggia, J., 2011. Does access to external finance improve productivity? Evidence from a natural experiment. *J. Financ. Econ.* 99, 184–203.

Butler, A.W., Cornaggia, J., Gurun, U.G., 2016. Do local capital market conditions affect consumers' borrowing decisions? *Manag. Sci.* 63 (12), 4175–4187.

Celerier, C., Matray, A., 2016. Bank branch supply and the unbanked phenomenon. Working Paper, SSRN, 2392278

Cookson, J.A., 2010. Institutions and casinos: an empirical investigation of the location of Indian casinos. *J. Law Econ.* 53, 651–687.

Cookson, J.A., 2018. When saving is gambling. *J. Financ. Econ.* 129 (1), 24–45.

Cronqvist, H., Previtro, A., Siegel, S., White, R.E., 2016. The fetal origins hypothesis in finance: Prenatal environment, the gender gap, and investor behavior. *Rev. Financ. Stud.* 29 (3), 739.

Cronqvist, H., Siegel, S., 2015. The origins of savings behavior. *J. Polit. Econ.* 123 (1), 123–169.

Debbaut, P., Ghent, A., Kudlyak, M., 2016. The CARD Act and young borrowers: the effects and the affected. *J. Money Credit Bank.* 48 (7), 1495–1513.

Demirguc-Kunt, A., Klapper, L., 2013. Measuring financial inclusion: Explaining variation across and within countries. *Brook. Pap. Econ. Act.* 44 (1), Spring 2013

Detting, L.J., Hsu, J.W., 2014. Returning to the nest: debt and parental co-residence among young adults. Technical Report. FEDS Working Paper 2014-80. Board of Governors of the Federal Reserve System (U.S.)

Dimitrova-Grajzl, V., Grajzl, P., Guse, A.J., 2014. Jurisdiction, crime, and development: the impact of Public Law 280 in Indian country. *Law Soc. Rev.* 48 (1), 127–160.

Dimitrova-Grajzl, V., Grajzl, P., Guse, A.J., Todd, R.M., 2015. Consumer credit on American Indian reservations. *Econ. Syst.* 39 (3), 518–540.

Dimmock, S.G., Kouwenberg, R., Mitchell, O.S., Peijnenburg, K., 2016. Ambiguity aversion and household portfolio choice puzzles: empirical evidence. *J. Financ. Econ.* 119 (3), 559–577.

Dippel, C., 2014. Forced coexistence and economic development: evidence from American Indian reservations. *Econometrica* 82 (6), 2131–2165.

Dobbie, W., Goldsmith-Pinkham, P., Yang, C.S., 2017. Consumer bankruptcy and financial health. *Rev. Econ. Stat.* 99 (5), 853–869.

Dokko, J., Li, G., Hayes, J., 2015. Credit scores and committed relationships. Technical Report. FEDS Working Paper 2015-81 Board of Governors of the Federal Reserve System (U.S.)

Favara, G., Imbs, J., 2015. Credit supply and the price of housing. *Am. Econ. Rev.* 105 (3), 958–992.

Giannetti, M., Wang, T.Y., 2016. Corporate scandals and household stock market participation. *J. Finance* 71 (6), 2591–2636.

Goetz, M.R., Laeven, L., Levine, R., 2016. Does the geographic expansion of banks reduce risk? *J. Financ. Econ.* 120 (2), 346–362.

Goldberg-Ambrose, C., 1997. *Planting Tail Feathers: Tribal Survival and Public Law 280.* University of California, Los Angeles, CA.

Gopalan, R., Mukherjee, A., Singh, M., 2016. Do debt contract enforcement costs affect financing and asset structure? *Rev. Financ. Stud.* 29 (10), 2774–2813.

Grinblatt, M., Keloharju, M., Linnainmaa, J.T., 2012. IQ, trading behavior, and performance. *J. Financ. Econ.* 104 (2), 339–362.

Guiso, L., Sapienza, P., Zingales, L., 2004a. Does local financial development matter? *Q. J. Econ.* 94 (3), 526–556.

Guiso, L., Sapienza, P., Zingales, L., 2004b. The role of social capital in financial development. *Am. Econ. Rev.* 94 (3), 526–556.

Guiso, L., Sapienza, P., Zingales, L., 2008. Trusting the stock market. *J. Finance* 63 (6), 2557–2600.

Gurun, U.G., Stoffman, N., Yonker, S.E., 2017. Trust busting: the effect of fraud on investor behavior. *Rev. Financ. Stud.* 31 (4), 1341–1376.

Javaratne, J., Strahan, P.E., 1996. The finance-growth nexus: evidence from bank branch deregulation. *Q. J. Econ.* 111 (3), 929–969.

Jimenez, V., Song, S.C., 1998. Concurrent tribal and state jurisdiction under Public Law 280. *Am. Univ. Law Rev.* 47 (August), 1627–1707.

Karlan, D., Zinman, J., 2010. Expanding credit access: using randomized supply decisions to estimate the impacts. *Rev. Financ. Stud.* 23 (1), 433–464.

Karpoff, J.M., Rice, E.M., 1989. Organizational form, share transferability, and firm performance: evidence from the ANCSA corporations. *J. Financ. Econ.* 24, 69–105.

Keys, B.J., Wang, J., 2019. Minimum payments and debt payoff in consumer credit cards. *J. Financ. Econ.* 131 (3), 528–548.

King, R.G., Levine, R., 1993. Finance and growth: Schumpeter might be right. *Q. J. Econ.* 108, 717–737.

Knüpfper, S., Rantapuska, E.H., Sarvimäki, M., 2017. Formative experiences and portfolio choice: evidence from the Finnish great depression. *J. Finance* 72 (1), 133–166.

Krishnan, K., Nandy, D., Puri, M., 2015. Does financing spur small business productivity? Evidence from a natural experiment. *Rev. Financ. Stud.* 28 (6), 1768–1809.

Kuhnen, C.M., Miu, A.C., 2017. Socioeconomic status and learning from financial information. *J. Financ. Econ.* 124 (2), 349–372.

Levine, R., 2005. Finance and growth: theory and evidence. In: Aghion, P., Durlauf, S. (Eds.), *Handbook of Economic Growth*, 1. Elsevier, pp. 865–934. *Handbook of Economic Growth*, chapter 12

Levine, R., Loayza, N., Beck, T., 2000. Financial intermediation and growth: causality and causes. *J. Monet. Econ.* 46, 31–77.

²¹ For example, see “For the first time, more are mobile-banking than going to a branch,” Telis Demos, Wall Street Journal, January 12, 2016.

- Lusardi, A., Mitchell, O.S., 2011. Financial literacy around the world: an overview. *J. Pension Econ. Finance* 10, 497–508.
- Lusardi, A., Mitchell, O.S., 2014. The economic importance of financial literacy: theory and evidence. *J. Econ. Lit.* 52 (1), 5–44.
- Malmendier, U., Nagel, S., 2011. Depression babies: do macroeconomic experiences affect risk taking? *Q. J. Econ.* 126 (1), 373–416.
- Malmendier, U., Nagel, S., 2016. Learning from inflation experiences. *Q. J. Econ.* 131 (1), 53–87.
- Malmendier, U., Tate, G., Yan, J., 2011. Overconfidence and early-life experiences: the effect of managerial traits on corporate financial policies. *J. Finance* 66 (5), 1687–1733.
- McDevitt, R.C., Sojourner, A., 2016. Demand, regulation, and welfare on the margin of alternative financial services. Working Paper.
- Melzer, B.T., 2011. The real costs of credit access: evidence from the payday lending market. *Q. J. Econ.* 126 (1), 517–555.
- Osili, U.O., Paulson, A.L., 2008. Institutions and financial development: evidence from international migrants in the United States. *Rev. Econ. Stat.* 90 (3), 498–517.
- Parker, D.P., 2012. The effects of legal institutions on access to credit: evidence from American Indian reservations. Working Paper.
- Rice, T., Strahan, P.E., 2010. Does credit competition affect small-firm finance? *J. Finance* 65 (3), 861–889.
- Skrastins, J., 2017. Credit enforcement and output storage: evidence from grain warehouses. Working Paper, Washington University Olin School of Business.